## WORKING PAPERS

# Borders and Growth* 

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#### Abstract

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#### Abstract

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# Borders and Growth* 

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#### Abstract

This paper presents a framework to understand and measure the effects of political borders on economic growth and per capita income levels. In our model, political integration between two countries results in a positive country size effect and a negative effect through reduced openness vis-à-vis the rest of the world. Additional effects stem from possible changes in other growth determinants, besides country size and openness, when countries are merged. We estimate the growth effects that would have resulted from the hypothetical removal of national borders between pairs of adjacent countries under various scenarios. We identify country pairs where political integration would have been mutually beneficial. We find that full political integration would have slightly reduced an average country's growth rate, while most countries would benefit from a more limited form of merger, involving higher economic integration with their neighbors.


JEL Codes: F1, O5
Keywords: Economic integration, political unions, growth.

[^0]
## 1 Introduction

Some Questions. Are existing national borders good or bad for economic growth? What would the growth rate of per capita income in Canada have been if its border with the United States had not existed, that is, if they had been a single country? What if Italy and France had merged? In brief, what is the effect of borders on economic performance?

National borders constitute barriers to economic exchange, and may therefore reduce gains from specialization and trade. By contrast, removing national borders allows the formation of larger domestic markets, which may have a positive effect on productivity and growth if market size matters for economic activity. However, as stressed in the regional-integration literature, removing borders between regions while maintaining barriers with the rest of the world can bring about not only trade creation but also trade reduction. Moreover, national borders may shield some countries from slow-growing neighbors, and their removal might therefore reduce productivity and economic growth.

In order to estimate the effects of borders on growth one needs to answer three distinct but related questions:

1) Is openness good for growth?
2) Is a large market size good for growth?
3) Does a country's openness depend on its size?

While there exist extensive literatures addressing each of those questions separately, there are very few theoretical and empirical analyses that look at the effects of size and openness on growth jointly, treating openness as an endogenous function of size and other determinants. This paper argues that a simultaneous approach is essential. We will provide a unified exploration of those three important questions, and provide new quantitative answers.

We will then use our results to address whether existing national borders have been good or bad for growth. Specifically, we will ask a counterfactual question:
4) Would existing countries have gained much from merging with their neighbors?

In a nutshell, our answers to the first three questions will be "yes". Our answer to the last one is "it depends".

Summary of the Paper. In this paper we first present a simple theoretical framework accounting for the effects of openness and size on income and growth. In our model, politically integrated economies can save on trading costs, generating a market size effect of political integration. However, trade openness responds endogenously to political integration. All other things being equal, in a political economy equilibrium, larger countries tend to choose higher trade barriers with respect to the rest of the world. Therefore, political integration, by increasing the size of countries and hence their barriers, also generates a trade reduction effect. Finally, political integration can induce changes in the other determinants of steady-state income levels, besides country size, an effect we call the steady-state determination effect. Within our stylized framework, we derive closed-form solutions for the relationship between steady-state income per capita, openness and country size. We also obtain a relationship between openness, barriers to trade and size. We then discuss necessary and sufficient conditions for a positive effect of political mergers on steady-state income per capita and on economic growth.

In the second part of this paper we provide an empirical methodology to evaluate the effect of national borders on economic growth. We estimate the effect of market size on economic growth in a cross-country context. In our specification, derived directly from the model, market size can be increased by two means: expanding the internal market or gaining greater access to foreign markets. Consistent with our theoretical framework, growth is affected both by openness and domestic size, which also interact with a negative sign (the effect of a larger domestic size is reduced at higher levels of openness, and the effect of openness is smaller for a larger domestic size) Openness itself is estimated - simultaneously with growth - as an endogenous variable which is affected by domestic size, among other determinants. We find robust evidence of positive effects of openness and size on growth, and of a negative effect of size on openness.

Together, these estimates allow us to quantify the economic effect of specific borders by creating hypothetical merged countries (for example the one that would result from the United States merging with Canada or France merging with Italy), and estimating what their growth rate would have been over the sample period. This empirical exercise corresponds exactly to our theoretical counterfactual. We present estimates of the market size effect, the trade reduction effect and the steady-state determination effect for all pairs of adjacent coun-
tries and proximate islands for which data is available (123 hypothetical pairwise mergers). ${ }^{1}$ We identify cases where political integration would be mutually beneficial to the merging countries, and discuss the conditions under which this occurs.

We find that the complete removal of political borders between neighbors (full integration) would not be growth-enhancing on average. In fact, on average it would bring about a (slight) negative net effect on growth: -0.112 percentage points of annual growth. In other words, the typical country would lose from pairwise political integration with a neighbor. Moreover, out of 123 hypothetical mergers, only 14 would bring about benefits for both countries involved in the merger. By contrast, integration of domestic markets in which each country remains politically independent would tend, on average, to increase growth performances across countries. In summary, while we find a few cases in which countries could benefit economically from full political integration with their neighbors, a more promising avenue for most countries would be to extend the size of their markets by lowering barriers to trade with their neighbors and the rest of the world, while maintaining their political independence.

Literature. This paper builds on and contributes to several related literatures. There exists a vast theoretical and empirical literature on the relationship between national borders and trade. Recent important contributions that directly document the effect of national borders on trade include McCallum [1995] and Helliwell [1998]. ${ }^{2}$ This literature suggests that removing national borders would substantially reduce barriers to interregional trade. A second, related body of work in the field of international trade is the extensive literature on the effects of regional integration on trade, efficiency and welfare. This literature has stressed how removing a specific political border can result in trade reduction vis-à-vis third countries. In particular, the classical theory of customs unions has pointed out the welfare

[^1]losses from trade diversion. ${ }^{3}$ More generally, the literature has studied the costs and benefits of regional integration in a second-best world in which integrated markets face barriers with the rest of the world. ${ }^{4}$ A third, related body of work in the field of international trade, which is immediately relevant for our analysis, has focused on the endogenous formation of barriers to international trade. ${ }^{5}$

While we build on the theory and empirics of international trade, our approach in this paper is more closely linked to the growing macroeconomics literature on the relationship between openness, market size and growth. The relationship between openness and growth has been the focus of numerous studies, which include Sachs and Warner [1995], Frankel and Romer [1998], Rodríguez and Rodrik [2000], Wacziarg [2001], and Alcalá and Ciccone [2001] among others. In this literature the effect of openness on economic performance is usually studied without controlling for countries' domestic size. A second, smaller but growing literature has focused on the importance of market size for productivity and growth. In particular, our paper is most closely related to Alesina, Spolaore and Wacziarg [2000, henceforth ASW], who have stressed that a) the effects of market size and openness on economic performance should be studied jointly, and b) openness should have a larger effect for smaller countries, while domestic size should have a larger effect for closed countries. ${ }^{6}$

Our paper builds on the ASW framework. However, it addresses a different set of issues and, consequently, differs from the ASW framework in several key respects. A central difference with the ASW framework is that we treat openness and barriers to trade as fully endogenous, and we consequently model and estimate growth and openness simultaneously. The ASW analysis focused on how exogenous changes in the level of trade barriers affect the number and size of nations in a world of endogenous borders. By contrast, our focus is on the effects of (counterfactual) changes in the configuration of borders on a country's level of openness and economic performance. Hence, in our empirical framework we jointly estimate

[^2]the effects of market size and openness on growth and income levels and the relationship between openness and market size. Moreover, we use our estimates of the growth effects of market size and openness to construct the empirical analog to our theoretical counterfactuals, and to estimate the effects of specific borders on growth and income levels. Thus, we view this paper as providing a novel way to examine the relationship among market size, openness and growth, and to provide quantitative estimates of the economic effects of national borders.

Outline. This paper is structured as follows: Section 2 presents a model of economic growth based on scale effects, and analyzes the effect of borders on growth in this context. Section 3 describes our empirical methodology for estimating the border effect and discusses extensions. Section 4 presents our empirical results, and Section 5 concludes.

## 2 A Model of Political Integration and Growth

### 2.1 Assumptions of the Model

This section presents a stylized model that links political borders, international openness and productive activity. In this model, market size affects growth and income levels, and depends both on the degree of openness of the economy and on country size. ${ }^{7}$ Openness, measured by the ratio of trade to output, is itself endogenous, and responds to country size via endogenous barriers to trade.

There is a continuum of regions, measured on the interval $[0, W]$. Time is continuous. The intertemporal utility function in each region $i$ is given by:

$$
\begin{equation*}
U_{i}=\int_{0}^{\infty} \ln c_{i}(t) e^{-\rho t} d t \tag{1}
\end{equation*}
$$

where $c_{i}(t)$ denotes consumption at time $t$ by the representative household living in region $i$, and $\rho>0 .{ }^{8}$ At time $t$ region $i$ 's capital and labor are denoted, respectively, by $K_{i}(t)$ and $L_{i}(t)$. Both inputs are supplied inelastically and are not mobile across regions. Each region $i$ produces a specific intermediate input $X_{i}(t)$ using the region-specific capital according to

[^3]the following linear production function:
\[

$$
\begin{equation*}
X_{i}(t)=K_{i}(t) \tag{2}
\end{equation*}
$$

\]

There exists a unique final good. Each region $i$ produces $y_{i}(t)$ units of the final good, according to the production function:

$$
\begin{equation*}
y_{i}(t)=\left(\int_{0}^{W} x_{j i}^{\alpha}(t) d j\right) L_{i}^{1-\alpha}(t) \tag{3}
\end{equation*}
$$

with $0<\alpha<1$. $x_{j i}(t)$ denotes the amount of intermediate input $j$ used in region $i$ at time $t$.
Regions are divided into $N$ countries. Country 1 includes all regions in the interval [0, $\left.S_{1}\right]$; country 2 includes all regions in the interval $\left[S_{1}, S_{1}+S_{2}\right.$ ], country $n$ includes all regions in the interval [ $S_{n-1}, S_{n-1}+S_{n}$ ], etc. Each region inelastically supplies one unit of labor (i.e., $L_{i}(t)=1$ for every $i$ at every $t$.). Hence, the "size" of country 1 (measured by total labor) is equal to $S_{1}$, the size of country 2 is $S_{2}$, the size of country $n$ is $S_{n}$, etc.

Intermediate inputs can be traded across regions that belong to the same country at no cost (i.e., we assume no internal barriers to trade). By contrast, if one unit of an intermediate good $j$ that belongs to country $a$ is shipped to a region that belongs to a different country (say, country $b$ ), only ( $1-\xi_{a}-\xi_{b}$ ) units of the intermediate good will arrive, where $0<\xi_{a}+\xi_{b} \leq 1$. Hence, the levels of $\xi_{n}$ 's measure barriers to trade across national borders.

### 2.2 Equilibrium

Intermediate inputs are sold in perfectly competitive markets. In equilibrium, each unit of each input will be sold at a price equal to its marginal product. All regions that belong to the same country will use identical levels of a given input. Hence, we can let $x_{i n}$ denote the amount of input $i$ used in each region of country $n$. Let $P_{i}(t)$ denote the market price of intermediate input $i$, where region $i$ belongs to country $a$. Therefore, for every input $i$ belonging to a country $a$ and for every country $n \neq a$ we must have:

$$
\begin{equation*}
P_{i}(t)=\alpha x_{i a}^{\alpha-1}(t)=\alpha\left(1-\xi_{a}-\xi_{n}\right)^{\alpha} x_{i n}^{\alpha-1}(t) \tag{4}
\end{equation*}
$$

By using the above equation (4) and the resource constraint, as shown in Appendix 1, we can obtain the equilibrium price of each input $i$ produced in country $a$ :

$$
\begin{equation*}
P_{i}(t)=\alpha\left[S_{a}+\sum_{n \neq a} S_{n}\left(1-\xi_{a}-\xi_{n}\right)^{\frac{\alpha}{1-\alpha}}\right]^{1-\alpha} K_{i}(t)^{\alpha-1} \tag{5}
\end{equation*}
$$

Households' net assets in region $i$ are identical to the stock of region-specific capital $K_{i}(t)$. Since each unit of capital yields one unit of intermediate input $i$, the net return to capital is equal to the market price of intermediate input $P_{i}(t)$ (for simplicity, we assume no depreciation). From standard intertemporal optimization we have the following Euler equation for consumption in region $i$ belonging to country $a$ :

$$
\begin{equation*}
\frac{d c_{i t}}{d t} \frac{1}{c_{i t}}=P_{i}(t)-\rho=\alpha\left[S_{a}+\sum_{n \neq a} S_{n}\left(1-\xi_{a}-\xi_{n}\right)^{\frac{\alpha}{1-\alpha}}\right]^{1-\alpha} K_{i}(t)^{\alpha-1}-\rho \tag{6}
\end{equation*}
$$

Hence, the steady-state level of capital in each region $i$ belonging to country $a$ is:

$$
\begin{equation*}
K_{i}^{s s}=\left(\frac{\alpha}{\rho}\right)^{\frac{\alpha}{1-\alpha}}\left[S_{a}+\sum_{n \neq a} S_{n}\left(1-\xi_{a}-\xi_{n}\right)^{\left.\frac{\alpha}{1-\alpha}\right]}\right. \tag{7}
\end{equation*}
$$

The steady-state level of output per capita in a region $i$ of a country of size $S_{a}$ is given by: ${ }^{9}$

$$
\begin{equation*}
y_{i}^{s s}=\left(\frac{\alpha}{\rho}\right)^{\frac{\alpha}{1-\alpha}}\left[S_{a}+\sum_{n \neq a} S_{n}\left(1-\xi_{a}-\xi_{n}\right)^{\frac{\alpha}{1-\alpha}}\right] \tag{8}
\end{equation*}
$$

Our model has standard neoclassical implications as far as the growth rate is concerned. In particular, at each point in time the growth rate of income per capita is positively related to steady-state income per capita and negatively related to the current (initial) level of income: ${ }^{10}$

$$
\begin{equation*}
\frac{d \ln y_{n}(t)}{d t}=f\left(y_{n}^{s s}, \quad y_{n, t-\tau}\right) \tag{9}
\end{equation*}
$$

with

$$
\begin{equation*}
\frac{\partial f}{\partial y_{n}^{s s}}>0 \quad, \quad \frac{\partial f}{\partial y_{n, t-\tau}}<0 \tag{10}
\end{equation*}
$$

Therefore, the effects of size, openness or other variables on the level of income per capita also translate into effects on the growth rate in the transition to the steady-state. Thus, in this theoretical section we will focus our analysis on steady-state income. Implications for growth will be studied in the empirical section.

### 2.3 Steady-state income, country size and openness

We are now ready to derive the relationship between income per capita, country size and openness to trade. Let $O_{a}$ measure the exports to output ratio in country $a$. We will refer

[^4]to this measure as "openness". ${ }^{11}$ Steady-state $O_{a}$ can be easily derived as follows. Each region in country $a$ will use $x_{i a}^{s s}$ units of inputs locally, and will sell an equal amount $x_{i a}^{s s}$ to each of the other $S_{a}-1$ regions belonging to country $a$. Hence, total exports of input $i$ will be given by $K_{i}^{s s}-S_{a} x_{i a}^{s s}$. Since all regions in country $a$ export the same amount, total exports in country $a$ are given by $\left(K_{i}^{s s}-S_{a} x_{i a}^{s s}\right) S_{a}$. Country $a$ 's total output is given by $y_{i}^{s s} S_{a}$. Therefore, the exports to output ratio $O_{a}$ in steady-state is given as follows:
\[

$$
\begin{equation*}
O_{a}^{s s}=\frac{\left(K_{i}^{s s}-S_{a} x_{i a}^{s s}\right) S_{a}}{y_{i}^{s s} S_{a}} \tag{11}
\end{equation*}
$$

\]

By using the expression for $x_{i a}$ derived in Appendix 1 (equation 51) and equations (7) and (8), we can write the equilibrium steady-state level of openness as follows:

$$
\begin{equation*}
O_{a}^{s s}=\frac{\sum_{n \neq a} S_{n}\left(1-\xi_{a}-\xi_{n}\right)^{\frac{\alpha}{1-\alpha}}}{S_{a}+\sum_{n \neq a} S_{n}\left(1-\xi_{a}-\xi_{n}\right)^{\frac{\alpha}{1-\alpha}}} \tag{12}
\end{equation*}
$$

$O_{a}^{s s}$ is decreasing in $\xi_{a}$ and in $S_{a}$ :

Proposition 1 - Openness in steady-state is inversely related to a country's size and to a country's barriers to trade.

Equation (12) can be used to express steady-state output per capita in equation (8) as a function of a country's size and openness:

$$
\begin{equation*}
y_{i}^{s s}=\left(\frac{\alpha}{\rho}\right)^{\frac{\alpha}{1-\alpha}} \frac{S_{a}}{1-O_{a}^{s s}} \tag{13}
\end{equation*}
$$

As equation (12) clearly shows, $O_{a}^{s s}$ itself is a function of size $S_{a}$. However, it is useful to consider the partial effects of size and openness on income per capita and their interaction. That is, it is useful to consider the effect of size on income for given openness and the effect of openness of income for a given size. Specifically, we have:

Proposition 2: Income per capita in steady-state is increasing in country size (for given openness) and increasing in openness (for given country size). The positive effect of size is higher the lower is openness, while the positive effect of openness is higher the smaller is size.

[^5]Formally,

$$
\begin{gather*}
\frac{\partial y_{i}^{s s}}{\partial S_{a}}=\left(\frac{\alpha}{\rho}\right)^{\frac{\alpha}{1-\alpha}} \frac{1}{1-O_{a}^{s s}}>0  \tag{14}\\
\frac{\partial y_{i}^{s s}}{\partial O_{a}^{s s}}=\left(\frac{\alpha}{\rho}\right)^{\frac{\alpha}{1-\alpha}} \frac{S_{a}}{\left(1-O_{a}^{s s}\right)^{2}}>0  \tag{15}\\
\frac{\partial\left(y_{i}^{s s}\right)^{2}}{\partial S_{a} \partial O_{a}^{s s}}=-\left(\frac{\alpha}{\rho}\right)^{\frac{\alpha}{1-\alpha}} \frac{1}{\left(1-O_{a}^{s s}\right)^{2}}<0 \tag{16}
\end{gather*}
$$

As we will see in the empirical section, the data are consistent with Proposition 2's main insights, that is, a) positive effects of size and openness on income per capita in steady-state (and hence growth in the transition to the steady-state), and b) a negative "interaction" between size and openness - meaning that the effect of size is smaller for more open countries, and the effect of openness is smaller for larger countries.

As we have already mentioned, "openness" is an endogenous variable, and, even for given barriers, it does depend on size $S_{a}$. Moreover, as we will see below, barriers to trade should also be viewed as an endogenous function of size - a relationship that introduces an additional channel through which size can affect openness. These endogenous links between openness and size will be taken into account in the empirical analysis.

### 2.4 Endogenous Barriers to Trade

So far we have considered barriers to trade as given. We will now extend the analysis to allow for an endogenous determination of barriers. Specifically, we will assume that, for each country $n$, barriers are given as follows:

$$
\begin{equation*}
\xi_{n}(t)=\frac{\xi}{2}-\lambda_{n}(t) \tag{17}
\end{equation*}
$$

where $\lambda_{n}(t)$ is the endogenous reduction in barriers by country $n$ at time $t$.
It is reasonable to assume that lowering trade barriers entails political and administrative costs. We capture the costs of reducing one's barriers in a stylized manner, by assuming a convex cost of barriers reduction:

$$
\begin{equation*}
B_{n}(t)=\frac{\phi_{n}}{2}\left[\lambda_{n}(t)\right]^{2} \tag{18}
\end{equation*}
$$

On the other hand, trade barriers may bring about political benefits (rents, etc.) to a country's policy-makers. We will capture those rents as a simple, linear function of the
barriers. Specifically, we will assume that at each time $t$ the government of country $n$ will choose its barrier reduction $\lambda_{n}(t)$ in order to solve: ${ }^{12}$

$$
\begin{equation*}
\max _{\lambda_{n}(t)} \psi_{n} \xi_{n}(t)+\left(1-\psi_{n}\right) c_{n}(t)-B_{n}(t) \tag{19}
\end{equation*}
$$

$\psi_{n}$ is the weight given to the "political" benefits or rents associated with barriers to trade, while $1-\psi_{n}$ is the weight given to the consumption level of the representative consumer at time $t .{ }^{13}$

In general, at each point in time barriers will be a function of the political parameters $\psi_{n}$ and $\phi_{n}$ and of the determinants of the equilibrium consumption path. In particular, the steady-state level of barriers $\xi_{n}^{s s}$ will be given by the solution of the following optimization:

$$
\begin{equation*}
\max _{\lambda_{n}}\left\{\psi_{n} \xi_{n}+\left(1-\psi_{n}\right)\left(\frac{\alpha}{\rho}\right)^{\frac{\alpha}{1-\alpha}} c_{n}^{s s}-\frac{\phi_{n}}{2} \lambda_{n}^{2}\right\} \tag{20}
\end{equation*}
$$

Since steady-state consumption is equal to steady-state income in our model, we can substitute $c_{n}^{s s}$ in the above equation (20) with $y_{n}^{s s}$ from equation (8). Hence the steady-state level of barrier reduction $\lambda_{a}^{s s}$ for a country of size $S_{a}$ will be given by

$$
\begin{align*}
\lambda_{a}^{s s}= & \arg \max _{\lambda_{a}}\left\{\psi_{a}\left[\frac{\xi}{2}-\lambda_{a}\right]+\right.  \tag{21}\\
& \left.\left(1-\psi_{a}\right)\left(\frac{\alpha}{\rho}\right)^{\frac{\alpha}{1-\alpha}}\left[S_{a}+\sum_{n \neq a} S_{n}\left(1-\frac{\xi}{2}+\lambda_{n}-\xi_{n}\right)^{\frac{\alpha}{1-\alpha}}\right]-\frac{\phi_{a}}{2} \lambda_{a}^{2}\right\}
\end{align*}
$$

For each country $a=1,2, \ldots N$, the first-order condition for $\lambda_{a}^{*}$ is given as follows:

$$
\begin{equation*}
-\psi_{a}+\left(1-\psi_{a}\right) \frac{\alpha}{1-\alpha}\left(\frac{\alpha}{\rho}\right)^{\frac{\alpha}{1-\alpha}} \sum_{n \neq a} S_{n}\left(1-\frac{\xi}{2}+\lambda_{a}^{s s}-\xi_{n}\right)^{\frac{2 \alpha-1}{1-\alpha}}-\phi_{a} \lambda_{a}^{s s}=0 \tag{22}
\end{equation*}
$$

[^6]In general, the equilibrium level of barriers reduction in each country is a function of the size distribution of all countries. ${ }^{14}$ Other things being equal, smaller countries tend to have lower barriers. For example, in a world of two countries $\left(W=S_{a}+S_{b}\right)$ with identical political parameters $\left(\psi_{a}=\psi_{b}=\psi\right.$ and $\left.\phi_{a}=\phi_{b}=\phi\right)$ we have: ${ }^{15}$

$$
\begin{equation*}
\frac{d \lambda_{a}^{s s}}{d S_{a}}=-\frac{1-\psi}{\phi} \frac{\alpha}{1-\alpha}\left(\frac{\alpha}{\rho}\right)^{\frac{\alpha}{1-\alpha}}\left(1-\xi+\lambda_{a}^{s s}+\lambda_{b}^{s s}\right)^{\frac{2 \alpha-1}{1-\alpha}}<0 \tag{23}
\end{equation*}
$$

A simple closed-form solution can be obtained for the case $\alpha=1 / 2$. Then the degree of barrier reduction $\lambda_{a}^{*}$ that maximizes output per capita minus barriers reduction costs for a country of size $S_{a}$ is:

$$
\begin{equation*}
\lambda_{a}^{s s}=\frac{\left(1-\psi_{a}\right)\left(W-S_{a}\right)}{2 \phi_{a} \rho}-\frac{\psi_{a}}{\phi_{a}} \tag{24}
\end{equation*}
$$

which, again, implies a negative relationship between barrier reduction and size: ${ }^{16}$

$$
\begin{equation*}
\frac{d \lambda_{a}^{s s}}{d S_{a}}=-\frac{1-\psi_{a}}{2 \phi_{a} \rho}<0 \tag{25}
\end{equation*}
$$

Hence, we have the following:

Proposition 3 - All other things equal, larger countries will have less open trade policies - that is, they will choose smaller reductions of barriers $\left(\lambda_{a}^{s s}\right)$ - and, consequently, higher barriers $\xi_{a}^{s s}$.

Countries with lower costs of reduction $\left(\phi_{a}\right)$, lower weight on political rents $\left(\psi_{a}\right)$ or a lower discount rates $(\rho)$ will be more open (that is, will have a higher $\lambda_{a}$ ).

### 2.5 Political Mergers

Now, let us consider a merger between country $a$ (of size $S_{a}$ ) and country $b$ (of size $S_{b}$ ). To keep things simple we will assume $\alpha=1 / 2$ and $\psi_{a}=\psi_{b}=\psi$ and $\phi_{a}=\phi_{b}=\phi$. The steady-state levels of income per capita in country $a$ is:

$$
\begin{equation*}
y_{a}^{s s}=\left(\frac{1}{2 \rho}\right)\left[S_{a}+\sum_{n \neq a} S_{n}\left(1-\xi_{a}^{s s}-\xi_{n}^{s s}\right)\right] \tag{26}
\end{equation*}
$$

[^7]The steady-state income per capita in the new country of size $S_{m}=S_{a}+S_{b}$ will be:

$$
\begin{equation*}
y_{m}^{s s}=\left(\frac{1}{2 \rho}\right)\left[S_{m}+\sum_{n \neq m} S_{n}\left(1-\xi_{m}^{s s}-\xi_{n}^{s s}\right)\right] \tag{27}
\end{equation*}
$$

The net change in steady-state income for country $a$ will be given by:

$$
\begin{equation*}
y_{m}^{s s}-y_{a}^{s s}=\left(\frac{1}{2 \rho}\right)\left[S_{b}\left(\xi_{a}^{s s}+\xi_{b}^{s s}\right)-\left(W-S_{a}-S_{b}\right)\left(\xi_{m}^{s s}-\xi_{a}^{s s}\right)\right] \tag{28}
\end{equation*}
$$

In equation (28), the first term, $(1 / 2 \rho) S_{b}\left(\xi_{a}^{s s}+\xi_{b}^{s s}\right)$, measures the direct positive scale effect of the merger, which we call the market size effect. It is evaluated at the level of trade barriers prevailing before the merger and corresponds to adding the size of country $b$ to country $a$.

The second term in equation (28), $(1 / 2 \rho)\left(W-S_{a}-S_{b}\right)\left(\xi_{m}^{s s}-\xi_{a}^{s s}\right)$, measures the indirect negative effect of the merger, via a fall in openness. We call this effect the trade reduction effect. It corresponds to the increase in trade barriers between the regions of former country $a$ and the rest of the world (i.e., all other countries except country b), brought forth by the larger size of the merged country $\left(S_{m}\right)$. That is, this effect is due to the fact that the larger country will be less open with respect to the rest of the world.

Note that there is no guarantee that the net gain in terms of steady-state income (and growth) will be positive. That is, there is no guarantee that steady-state income per capita in the new, larger country will be higher than in country $a$ - i.e., that $y_{m}^{s s}-y_{a}^{s s}>0$.

From equation (24) we have:

$$
\begin{align*}
& \xi_{a}^{s s}=\frac{\xi}{2}-\lambda_{a}^{s s}=\frac{\xi}{2}-\frac{(1-\psi)\left(W-S_{a}\right)}{2 \phi \rho}+\frac{\psi}{\phi}  \tag{29}\\
& \xi_{b}^{s s}=\frac{\xi}{2}-\lambda_{b}^{s s}=\frac{\xi}{2}-\frac{(1-\psi)\left(W-S_{b}\right)}{2 \phi \rho}+\frac{\psi}{\phi}  \tag{30}\\
& \xi_{m}^{s s}=\frac{\xi}{2}-\lambda_{m}^{s s}=\frac{\xi}{2}-\frac{(1-\psi)\left(W-S_{m}\right)}{2 \phi \rho}+\frac{\psi}{\phi} \tag{31}
\end{align*}
$$

which, when substituted in equation (28), imply the following:

Proposition 4-A necessary and sufficient condition for $y_{m}^{s s}-y_{a}^{s s}>0$ is:

$$
\begin{equation*}
S_{m}=S_{a}+S_{b}>\frac{3(1-\psi) W-2 \rho(\xi \phi+2 \psi)}{2(1-\psi)} \tag{32}
\end{equation*}
$$

The intuition for this results is as follows. A higher $S_{m}$ means a bigger positive effect from the merger via the market size effect, because the two merging countries had larger barriers
between themselves before the merger. A larger $S_{m}$ (relative to $W$ ) also means that the rest of the world is relatively smaller, and therefore the openness reduction effect (with respect to the rest of the world) has smaller costs.

It is important to notice that even if a merger increases income per capita, it does not necessarily imply an increase in consumption per capita and welfare. In order to calculate changes in consumption and welfare one should subtract the costs related to barriers reduction and any other costs associated with a merger. For example, a merger may bring about direct costs in order to eliminate internal barriers to trade. A merger may also imply higher "heterogeneity" costs due to different preferences over public goods, more costly coordination, etc. ${ }^{17}$ In our empirical exercises we will focus on changes of income per capita.

### 2.6 Other Determinants of Steady-State Income Levels

In our model so far, different countries' steady-states differ only because 1) their size differs and 2) as a result, their level of openness also differs. There are obviously many other differences across countries, apart from size, that could yield differences in steady-state income levels and openness. In the context of our model, the $\psi, \phi$ and $\rho$ parameters could differ across individual countries. Particularly patient countries, or countries where the costs of openness reduction are lower (for example through natural access to the sea, proximity to trading partners, and other geographic factors) will have higher levels of steady-state income and greater levels of openness, all else equal.

Such differences will not affect country $a$ 's growth performance under political integration with country $b$, unless they affect the other determinants of steady-state income levels and openness within country $a$. But it is easy to see that a merger between country $a$ and country $b$, when they differ along these other dimensions, will change the growth effect of the merger on country $a$, to the extent that the merger affects these parameters within country $a$. We should stress again that this would only occur if country $a$ 's steady-state and openness determinants (other than its size and induced openness level) would change under political integration. This could occur as the result of factor movements such as migration or capital flows, or changes in geographic factors brought forth by the removal of borders. ${ }^{18}$

[^8]In the case where countries differ in $\psi, \phi$ and $\rho$, the thought experiment described above to evaluate the growth incidence of political mergers can be amended to account for changes in steady-state determinants under a merger. For example, if countries have different costs $\phi$ 's, the analysis can be easily generalized as follows. Let $\phi_{m}$ denote the costs of barriers reduction in the unified country of size $S_{m}$. Then we have the following:

Proposition 5: A political merger between a country of size $S_{a}$ and a country of size $S_{b}$ will increase income in country a in steady state (that is, $y_{m}^{s s}-y_{a}^{s s}>0$ ) if and only if the following condition holds:

$$
\begin{align*}
& S_{b}\left[\xi+\frac{\psi}{\phi_{a}}+\frac{\psi}{\phi_{b}}-\frac{(1-\psi)\left(W-S_{a}\right)}{\rho \phi_{a}}\right.  \tag{33}\\
& \left.-\frac{(1-\psi)\left(W-S_{b}\right)}{\rho \phi_{b}}\right] \\
> & \left(W-S_{a}-S_{b}\right)\left[\frac{\psi}{\phi_{m}}-\frac{(1-\psi)\left(W-S_{a}-S_{b}\right)}{\rho \phi_{m}}-\frac{\psi}{\phi_{a}}\right. \\
& \left.+\frac{(1-\psi)\left(W-S_{a}\right)}{\rho \phi_{a}}\right]
\end{align*}
$$

In what follows, we will label the effect of potential changes in steady-state determinants, besides openness and country size, as the steady-state determination effect.

## 3 Estimating the Growth Effect of Borders

### 3.1 Basic Methodology

The model presented above, specifically Proposition 2, suggests that income in steady-state is positively related to both country size and openness, and negatively related to their interaction. Hence, growth in the transition to the steady-state will also be a function of such variables. A specification consistent with those insights is:

$$
\begin{align*}
\log \frac{y_{a t}}{y_{a t-\tau}}= & \beta_{0}+\beta_{1} \log y_{a t-\tau}+\beta_{2} O_{a t}+\beta_{3} \log S_{a t} \\
& +\beta_{4} O_{a t} \log S_{a t}+\beta_{5}^{\prime} Z_{a t}+\varepsilon_{a t} \tag{34}
\end{align*}
$$

where $a$ refers to a country, $S_{a t}$ denotes country size, $O_{a t}$ denotes trade openness, $y_{a t}$ denotes per capita income, and $Z_{a t}$ is a vector of control variables. Compared to our model, we have
merger. This could affect the level of openness of the country, and consequently its growth rate. See for instance Sachs [2001], and Gallup et al. [1999] for evidence on the importance of geography for growth.
simply added additional determinants of steady-state income levels (the $Z_{a t}$ variables), which the model abstracts from, and an error term. The predictions of our model are that $\beta_{2}>0$, $\beta_{3}>0$ and $\beta_{4}<0$.

In our model, Propositions 1 and 3 suggests that openness is negatively related to country size. The second part of our econometric model reflects the negative relationship between trade openness and country size:

$$
\begin{equation*}
O_{a t}=\alpha_{0}+\alpha_{1} \log S_{a t}+\alpha_{2}^{\prime} W_{a t}+\nu_{a t} \tag{35}
\end{equation*}
$$

where $W_{a t}$ is a vector of additional determinants of trade openness and the model predicts $\alpha_{1}<0$. In this econometric model, the exogenous variables are $S_{a t}, Z_{a t}$ and $W_{a t}$. We are considering the growth effect of an exogenous change in a country's size brought about by merging with a neighbor. Substituting equation (35) into (34), we obtain:

$$
\begin{align*}
\log \frac{y_{a t}}{y_{a t-\tau}}= & \gamma_{0}+\gamma_{1} \log y_{a t-\tau}+\gamma_{2} \log S_{a t}+\gamma_{3}\left(\log S_{a t}\right)^{2}+\gamma_{4}^{\prime} W_{a t} \log S_{a t} \\
& +\gamma_{5} \nu_{a t} \log S_{a t}+\gamma_{6}^{\prime} W_{a t}+\gamma_{7}^{\prime} Z_{a t}+\mu_{a t} \tag{36}
\end{align*}
$$

where the $\gamma$ coefficients are functions of the parameters of the growth and trade equations, as defined in Appendix 2.

Define $\Delta G_{a b t}$ as the change in growth of country $i$ resulting from its merger with country $b$. Since the only exogenous variable that has changed under a merger is country size, we term this particular exercise a "size merger". ${ }^{19}$ We focus on the expected effect on growth, as we have little knowledge of what the random component of growth or openness (captured by $\varepsilon_{a t}$ and $\nu_{a t}$ ) would have been had the countries been politically merged during the sample period. ${ }^{20}$ Assuming $E\left(\nu_{a t} \mid S_{a t}, S_{m t}, W_{a t}\right)=0$, the expected effect on the growth rate of country $a$ of merging with neighbor $b$, where the size of the merged country is denoted $S_{m t}$ $\left(=S_{a t}+S_{b t}\right)$, is:

$$
\begin{align*}
\Delta G_{a b t} & \equiv E\left(\left.\Delta \log \frac{y_{a t}}{y_{a t-\tau}} \right\rvert\, S_{a t}, S_{m t}, W_{a t}\right) \\
& =\log \left(\frac{S_{m t}}{S_{a t}}\right)\left[\gamma_{2}+\gamma_{3} \log \left(S_{a t} S_{m t}\right)+\gamma_{4}^{\prime} W_{a t}\right] \tag{37}
\end{align*}
$$

[^9]Thus, the effect of the merger on growth is a multiple of the percentage increase in country size, where the multiplicative factor depends on the determinants of openness, the estimated parameters of the model and the sizes of countries $a$ and $b$. Since our model predicts that $\gamma_{3}=\beta_{4} \alpha_{1}$ is positive, Proposition 3 is also directly apparent in equation (37).

In this basic setup, the induced effect of political integration on growth will depend on the home country's size, the size of the country it is considering merging with, and the determinants of the home country's trade openness volume. This combines three distinct effects of political integration on growth. Firstly, the direct (positive) effect of an increase in country size, equal to $\beta_{3}$ times the percentage increase in country size resulting from the merger $\left(\log \left(S_{m t} / S_{a t}\right)\right)$. Secondly, the indirect (negative) effect through openness reduction, which is equal to $\beta_{2} \alpha_{1}$ times the percentage increase in country size. Thirdly, the effect going through the interaction term, which captures the increasing impact of country size on growth as openness decreases. This effect, of ambiguous sign, depends on the determinants of $a$ 's openness level and the sizes of both $a$ and $b$, and is equal to $\beta_{4}\left(\alpha_{0}+\alpha_{1} \log \left(S_{m t} S_{a t}\right)+\alpha_{2}^{\prime} W_{a t}\right)$ times the percentage increase in country size. It should be noted that the determinants of openness $\left(W_{a t}\right)$ and the sizes of countries $a$ and $b$ can be such that the openness reducing effect of political integration outweighs the positive direct scale effect of merging. In this case, $\Delta G_{a b t}$ will be negative.

Finally, an exogenous change in openness yielding an equivalent expected change in economic growth without a political merger can be computed using equation (34) as:

$$
\begin{equation*}
E\left(\Delta O_{a t} \mid S_{a t}, S_{m t}, W_{a t}\right)=\frac{\Delta G_{a b t}}{\beta_{2}+\beta_{4} \log S_{a t}} \tag{38}
\end{equation*}
$$

The benefits of exogenous increases in openness can thus be directly compared to those of bilateral political mergers.

### 3.2 Changes in Conditioning Variables

Equations (37) and (38) implicitly assume that a political merger does not affect the determinants of the home country's steady-state income level, or the determinants of its openness levels, other than country size. For example, if France were to merge with Italy, France and Italy would each retain their own $Z_{a t}$ and $W_{a t}$ variables. These may include the savings rate, investment in human capital, characteristics of governance and government involvement in the economy, and gravity type factors such as geographic variables. As suggested in Section
2.6, this is clearly an extreme assumption since factors other than the size of the population alone would likely be different in each merged country under political integration, affecting both growth and the degree of openness. For example, increased migration and capital mobility across countries $a$ and $b$ under a merger will imply that the rates of investment in human and physical capital will differ compared to what they would have been in the separate countries. Taking this steady-state determination effect into consideration generates an additional sources of ambiguity in the sign of the overall effect of political integration on economic growth. Clearly, this effect would tend to be negative for the home country when the hypothetical merger is with a country with "worse" overall determinants of the steady-state income level than itself.

We can relax the assumption that political integration affects growth only through country size and the induced effect of changes in country size on trade openness by assuming that other conditioning variables will change in both merged units after political integration, and in particular that they will take on the same value in $a$ and $b$ under a merger. We term this alternative scenario "full integration".

There are obviously many ways to specify what values the other determinants of growth (the $Z$ variables) and openness (the $W$ variables) will take under full political integration. ${ }^{21}$ One reasonable assumption is that each of the merged countries would end up with the same population weighted average of the initial conditioning variables, which we can denote $Z_{m t}$ and $W_{m t}$, where the subscript $m$ denotes that a political merger has occurred and that the resulting variables are, where applicable, the population weighted averages of the regional measures. ${ }^{22}$ The resulting effect of a political merger on growth, $\Delta G_{a b t}^{m}$, is then computed

[^10]as:
\[

$$
\begin{align*}
\Delta G_{a b t}^{m} \equiv & E\left(\left.\Delta \log \frac{y_{a t}}{y_{a t-\tau}} \right\rvert\, S_{a t}, S_{m t}, W_{a t}, W_{m t}, Z_{a t}, Z_{m t}\right) \\
= & \log \left(\frac{S_{m t}}{S_{a t}}\right)\left(\gamma_{2}+\gamma_{3} \log \left(S_{a t} S_{m t}\right)+\gamma_{4}^{\prime} W_{a t}\right)  \tag{39}\\
& +\gamma_{1} \log \frac{y_{m t-\tau}}{y_{a t-\tau}}+\left[\gamma_{4}^{\prime} \log S_{m t}+\gamma_{6}^{\prime}\right]\left(W_{m t}-W_{a t}\right)+\gamma_{7}^{\prime}\left(Z_{m t}-Z_{a t}\right)
\end{align*}
$$
\]

This formulation includes the same size effects as equation (37), namely those that occur through the direct effect of market size, the indirect effect through trade reduction and the changes in the interaction term. But in addition to these effects, we now have the steady-state determination effect, equal to the terms in the second line of equation (39). ${ }^{23}$ An important consequence of this framework is that, under full political integration, expected growth will be equal for both country $a$ and country $b .^{24}$

To summarize, equations (37) and (39) result from two different assumptions about the effects of political integration on growth; one with complete averaging of steady-state determinants ("full integration"), the other with no changes in these variables ("size merger"). The effect of a hypothetical merger likely falls in between these two extremes. The corresponding estimates should therefore be viewed as extreme bounds on the effects of bilateral political mergers on economic growth.

### 3.3 Treatment of the Error Term

Above, we focused on estimating the expected effects of political mergers on growth, disregarding the unexplained portion of growth and openness in our counterfactual exercises. Whether to consider the residuals $\mu_{t}$ and $\nu_{t}$ from the growth and openness equations when evaluating the effects of borders on growth is largely a matter of interpretation. On the one hand, if one believes that they reflect omitted determinants of growth and openness, then

[^11]they should be treated as another steady-state determination variable (analogous to the $W$ and $Z$ variables). As it turns out, since the explained portion of growth and openness are typically on the order of $50 \%$ and $60 \%$ respectively, in our baseline regressions, accounting for the unexplained components of growth and openness could alter our estimates of the merger effects. On the other hand, if one believes that the residuals reflect true "randomness", then there is no good justification for including them in the analysis: we do not know what the random component of growth would have been, had the countries been merged over the sample period.

Since both interpretations seem equally defensible, we also present merger effects that take into account the estimated residuals. Fortunately, we can easily accommodate this change in our basic empirical methodology. Instead of computing the expected effect of a merger on growth, we can use:

$$
\begin{align*}
\Delta G_{a b t}^{e} & \equiv \Delta \log \frac{y_{a t}}{y_{a t-\tau}} \\
& =\log \left(\frac{S_{m t}}{S_{a t}}\right)\left[\gamma_{2}+\gamma_{3} \log \left(S_{a t} S_{m t}\right)+\gamma_{4}^{\prime} W_{a t}+\gamma_{5} \nu_{a t}\right] \tag{40}
\end{align*}
$$

and replace $\gamma_{2}, \gamma_{3}, \gamma_{4}$ and $\nu_{a t}$ with their regression estimates in computing the empirical $\Delta G_{a b t}^{e}$. In equation (40), the superscript " $e$ " indicates that the residual terms are taken into account. Note that since the error term of the growth regression, $\mu_{a t}$, is assumed to be unchanged between the merged and unmerged states, it gets differenced away from equation (40). ${ }^{25}$

Similarly, and perhaps more interestingly, we could treat the error terms as additional (unobserved) growth determinants, and compute the empirical $\Delta G_{a b t}^{m e}$ directly using the appropriate population weighted averages of the estimated residuals:

$$
\begin{align*}
\Delta G_{a b t}^{m e} \equiv & \Delta \log \frac{y_{a t}}{y_{a t-\tau}} \\
= & \log \left(\frac{S_{m t}}{S_{a t}}\right)\left[\gamma_{2}+\gamma_{3} \log \left(S_{a t} S_{m t}\right)+\gamma_{4}^{\prime} W_{a t}+\gamma_{5} \nu_{a t}\right] \\
& +\gamma_{1} \log \frac{y_{m t-\tau}}{y_{a t-\tau}}+\left[\gamma_{6}^{\prime}+\gamma_{4}^{\prime} \log S_{m t}\right]\left[W_{m t}-W_{a t}\right]  \tag{41}\\
& +\gamma_{7}^{\prime}\left[Z_{m t}-Z_{a t}\right]+\left(\nu_{m t}-\nu_{a t}\right)\left[\gamma_{5} \log S_{m t}\right]+\mu_{m t}-\mu_{a t}
\end{align*}
$$

[^12]where $\nu_{m t}$ and $\mu_{m t}$ are the population weighted averages of $\nu_{a t}$ and $\nu_{b t}$ and $\mu_{a t}$ and $\mu_{b t}$. Again, this equation involves the same terms as equation (40), with the steady-state determination effect (including that which results from merging the estimated unexplained portion of growth) added on.

### 3.4 Effects of Borders on Steady-State Income Levels

As explained in the theoretical section, because our model shares the dynamic features of the neoclassical growth model, it is straightforward to present our results in terms of steadystate income levels rather than growth. We do not observe steady-state income, but it can be estimated readily under the assumptions of our framework, because the right-hand side variables of equation (34) are the determinants of steady-state income levels. Theory delivers a growth equation of the following form, based on equation (9):

$$
\begin{equation*}
\log \frac{y_{a t}}{y_{a t-\tau}}=\lambda\left(\log y_{a}^{s s}-\log y_{a t-\tau}\right)+\varepsilon_{a t} \tag{42}
\end{equation*}
$$

where $y_{a t}$ is current income per capita, $y_{a t-\tau}$ is initial income per capita, and $y_{a}^{s s}$ is (unobserved) income in steady-state. ${ }^{26}$ Assume that the steady-state level of income takes the form:

$$
\begin{equation*}
\log y_{a}^{s s}=\delta_{1}+\delta_{2} O_{a t}+\delta_{3} \log S_{a t}+\delta_{4} O_{a t} \log S_{a t}+\delta_{5}^{\prime} Z_{a t} \tag{43}
\end{equation*}
$$

This specification choice for $\log y_{a}^{s s}$ reflects the fact that the right-hand side variables of empirical growth regressions (except initial income) are to be interpreted as the determinants of the steady-state level of income in the neoclassical growth model. On the other hand, our actual growth specification is that of equation (34):

$$
\begin{align*}
\log \frac{y_{a t}}{y_{a t-\tau}}= & \beta_{0}+\beta_{1} \log y_{a t-\tau}+\beta_{2} O_{a t}+\beta_{3} \log S_{a t} \\
& +\beta_{4} O_{a t} \log S_{a t}+\beta_{5}^{\prime} Z_{a t}+\varepsilon_{a t} \tag{44}
\end{align*}
$$

Substituting equation (43) into equation (42), we can write:

$$
\begin{align*}
\log \frac{y_{a t}}{y_{a t-\tau}}= & \lambda \delta_{1}+\lambda \delta_{2} O_{a t}+\lambda \delta_{3} \log S_{a t} \\
& +\lambda \delta_{4} O_{a t} \log S_{a t}+\lambda \delta_{5}^{\prime} Z_{a t}-\lambda \log y_{a t-\tau}+\varepsilon_{a t} \tag{45}
\end{align*}
$$

[^13]Thus, we can recover: ${ }^{27}$

$$
\begin{equation*}
\log y_{a}^{s s}=-\frac{\beta_{0}}{\beta_{1}}-\frac{\beta_{2}}{\beta_{1}} O_{a t}-\frac{\beta_{3}}{\beta_{1}} \log S_{a t}-\frac{\beta_{4}}{\beta_{1}} O_{a t} \log S_{a t}-\frac{1}{\beta_{1}} \beta_{5}^{\prime} Z_{a t} \tag{46}
\end{equation*}
$$

This provides a methodology for backing out the effects of political mergers on steady-state income levels. The percentage change in the steady-state income level of country after merging with country $b$ can be computed in terms of the reduced form parameters defined in Appendix 2, under the two scenarios under consideration - a pure size merger or full political integration:

$$
\begin{align*}
\Delta Y S S_{a b t} & \equiv E\left(\Delta \log y_{a}^{s s} \mid S_{a t}, S_{m t}, W_{a t}\right) \\
& =-\frac{1}{\gamma_{1}} \log \frac{S_{m t}}{S_{a t}}\left[\gamma_{2}+\gamma_{3} \log \left(S_{a t} S_{m t}\right)+\gamma_{4}^{\prime} W_{a t}\right] \tag{47}
\end{align*}
$$

and:

$$
\begin{align*}
\Delta Y S S_{a b t}^{m} \equiv & E\left(\Delta \log y_{a}^{s s} \mid S_{a t}, S_{m t}, W_{a t}, W_{m t}, Z_{a t}, Z_{m t}\right) \\
= & -\frac{1}{\gamma_{1}}\left[\log \frac{S_{m t}}{S_{a t}}\left(\gamma_{2}+\gamma_{3} \log \left(S_{a t} S_{m t}\right)+\gamma_{4}^{\prime} W_{a t}\right)\right.  \tag{48}\\
& \left.+\left(\gamma_{4}^{\prime} \log S_{m t}+\gamma_{6}^{\prime}\right)\left(W_{m t}-W_{a t}\right)+\gamma_{7}^{\prime}\left(Z_{m t}-Z_{a t}\right)\right]
\end{align*}
$$

Equations (47) and (48) are the analogs to equations (37) and (39), respectively, applied to income levels rather than growth. Note that equation (47) implies that $\Delta Y S S_{a b t}$ is simply $-1 / \gamma_{1}$ times $\Delta G_{a b t}$ - hence, since $\gamma_{1}=\beta_{1}$ is negative, the effect of a size merger on steadystate income will have the same sign as its effect on economic growth. However, the signs of $\Delta Y S S_{a b t}^{m}$ and $\Delta G_{a b t}^{m}$ may differ. This is because we have:

$$
\begin{equation*}
\Delta G_{a b t}^{m}=\gamma_{1}\left(\log \frac{y_{m t-\tau}}{y_{a t-\tau}}-\Delta Y S S_{a b t}^{m}\right) \tag{49}
\end{equation*}
$$

A country $a$ that has a positive steady-state level effect $\Delta Y S S_{a b t}^{m}$ of full integration may display a negative growth effect $\Delta G_{a b t}^{m}$ simply because it has a sufficiently low initial level of income relative to country $b$ (and hence enjoys relatively fast growth holding the steady-state level of income constant).

[^14]
## 4 Empirical Results

### 4.1 Estimates of the Growth and Openness Equations

### 4.1.1 Data and Estimators

Equations (34) and (35) can be readily estimated using cross-country data on growth, country size, openness and other control variables. Our measure of openness consists of the ratio of imports plus exports to GDP, a commonly used indicator of a country's overall level of openness. Moreover, this is precisely the measure derived in the theory of Section 2 (see in particular Propositions 2 and 3). The measure of country size consists of the log of a country's population. The $Z_{i t}$ variables are the common determinants of steady-state income levels in the cross-country literature: male and female human capital, the fertility rate, the ratio of government consumption to GDP and the rate of physical capital investment (see Barro and Sala-i-Martin [1995], chapter 12). Finally, the $W_{i t}$ variables consist of common determinants of openness such as geographic factors (land area, whether a country is landlocked or an island, whether it is an oil exporter) and the terms of trade shocks. In order to capture long-term phenomena, variables are averaged, where appropriate, over the sample period.

Our base estimates for calculating merger effects are based on PPP per capita income data from version 5.6 of the Penn World Tables (PWT). This 1960-1989 sample consists of 92 countries. Version 6.0 of this data has recently been circulated, extending the data to $1998 .{ }^{28}$ We use this data for the purpose of reestimating equations (34) and (35), as a robustness check. However, because some of the other conditioning variables are not as readily available for recent years, the updated sample only features 77 countries. Moreover, some "important" countries such as Germany are not part of this dataset for the entire sample period, precluding any calculation of the effect of political mergers on growth for such a key country in Europe. ${ }^{29}$ Therefore, in order to maximize the number of mergers we consider, and to base our estimates on the largest possible sample, we use estimates from version 5.6 of the PWT for the purpose of calculating merger effects. As shown below, the estimates of equations (34) and (35) do not differ much between versions of the PWT, so we are confident that using the more recent

[^15]data would not alter our results other than by limiting the country coverage.
One issue that arises clearly from our theoretical and empirical models is the endogeneity of openness (and the interaction term between openness and country size) in the growth equation. To address this, we treat equations (34) and (35) as a system of simultaneous equations to be estimated jointly. Our baseline results therefore consist of three-stage least squares estimates (3SLS). 3SLS treats all of the exogenous variables in the system (i.e. country size, initial per capita income, $Z_{i t}$ and $W_{i t}$ ) as potential instruments for the endogenous variables in the system (growth, openness and the interaction term between openness and country size). Given that openness and the interaction term are the only endogenous variables to appear on the right hand side of either equation in the system, only the $W_{i t}$ variables serve as instruments for them in the growth regression. As noted above, these variables consist of plausibly exogenous geographic and terms of trade variables. In addition to these instruments, we can gain precision by using additional instruments which do not necessarily appear as exogenous variables in either the trade or the growth equations. ${ }^{30}$ Finally, 3SLS allows for cross-equation covariance in the error terms $\varepsilon_{i t}$ and $\nu_{i t}$, generating potential efficiency gains. ${ }^{31}$ For the sake of robustness, we also present results obtained from seemingly unrelated regression (SUR), as well as regressions excluding the $Z_{i t}$ and $W_{i t}$ control variables.

### 4.1.2 Baseline Results

Tables 1 and 2 display results for the joint estimation of equation (34) and (35). The baseline estimates used for the merger calculations appear in column (1). The theoretical predictions are borne out empirically. Specifically, openness and country size are positively and significantly related to growth, while their interaction enters negatively and significantly. This is consistent with the model's results 1 and 3, and extends related findings in ASW [2000],

[^16]Ades and Glaeser [1999] and Alcalá and Ciccone [2003]. Moreover, as expected, country size affects openness negatively. This is consistent with our theoretical results 2 and 4, and extends previous findings in Alesina and Wacziarg [1998] and Wacziarg [2001].

Several additional observations are called for. First, the pattern of signs and statistical significance is unchanged when the $Z_{i t}$ and the $W_{i t}$ control variables are excluded from the system, and the magnitude of the coefficients of interest is raised. While this specification is likely to be tainted by omitted variables bias, it corresponds directly to the relationships derived from theory, where countries differed in no other way than size and openness. It is therefore reassuring that the predictions of the theory hold unconditionally as well as conditionally. Second, as in Frankel and Romer [1999], instrumenting for openness using geographic variables increases the magnitude of the estimated coefficient on trade openness compared to the specifications that do not account for the endogeneity of openness (SUR).

Since SUR estimates are tainted by endogeneity bias and since the unconditional estimates of column 2 are tainted by omitted variables bias, we rely on the estimates of column 1 as our benchmark to compute border effects. These estimates are not sensitive to small changes in the list of instruments or control variables. In fact, as argued in ASW and further shown here, the pattern and magnitude of coefficients on openness, country size and their interaction are remarkably robust, whether in cross-sectional or in panel (random effects) applications. As a consequence, it is also the case that our estimates of the effects of borders on growth and income levels are quite robust to changes in the specification.

Finally, Table 3 presents F-tests on the instruments, from simple OLS regressions of openness and the interaction term on all of the exogenous variables in the system, for the 1960-1989 data. These F-tests demonstrate that the instruments are jointly related to the variables they are instrumenting for, at high levels of statistical significance. Moreover, since our 3SLS estimates involve several instruments, exclusion restriction tests can be carried out for subsets of the instruments. Such Hausman tests showed that the instruments for openness and the interaction between openness and country size in the growth equation (namely the determinants of openness that are excluded from the growth equation) are indeed excludable from the regression. ${ }^{32}$

[^17]
### 4.1.3 Robustness

Our baseline results hold up when using the updated PWT 6.0 dataset for the period 19601998, despite the loss of 15 data points (Tables 1 and 2, columns 4-6). Due to this smaller sample, estimates are sometimes less statistically significant, but the pattern of signs and the magnitude of the coefficients are unchanged compared to the 1960-1989 dataset. Therefore, we are confident that our estimates of the border effects on growth would be qualitatively unchanged if we were to use coefficients from the updated dataset. As explained above, we refrain from using estimates obtained from the smaller dataset as this would result in a loss of 15 countries, in particular Germany.

The second robustness issue that we examine relates to our measure of openness. In an important paper, Alcalá and Ciccone (2001) have argued that commonly used volume measures of trade openness, obtained by taking the ratio of imports plus exports in exchange rate US dollars to GDP in exchange rate US dollar, may be inappropriate. The explanation is quite simple. Suppose that trade openness raises productivity, but does so more in the tradable than in the nontradable sector (a plausible assumption). This will lead to a rise in the relative price of nontradables, and a fall in conventionally measured openness under the assumptions that the demand for nontradables is relatively inelastic, because it may raise the denominator of the conventional measure of openness more than the numerator. So one may observe trade-induced productivity increases going hand in hand with a decline in conventional measures of openness. Alcalá and Ciccone propose an alternative measure, "real openness", defined as the ratio of imports plus exports in exchange rate US $\$$ to GDP in PPP US\$. This alternative measure will address the problem, since the denominator now corrects for international differences in the price of nontradable goods.

Tables 4 and 5 (columns 1 and 2) present 3SLS estimates of our baseline model using Alcalá and Ciccone's "real openness" measure, still using version 5.6 of PWT for the 196089 period. ${ }^{33}$ In both specifications with and without controls, our results on growth are confirmed and strengthened. Column 2 of Table 4 reveals an effect of openness on growth that
regression and suggested they were therefore valid instruments. Related to this, Frankel and Romer [1999] showed that an instrumental variable constructed from the geographic determinants of trade volumes were excludable from a regression of income levels on trade volumes.
${ }^{33}$ We thank Antonio Ciccone for providing us with the data.
has increased by $50 \%$ compared to the corresponding entry in Table 1. The magnitude and significance of the interaction term has also increased, while the magnitude of the coefficient on the $\log$ of population has decreased somewhat. In Table 5, while the effect of country size on openness still has a negative sign, its statistical significance and magnitude have fallen. Again, these results based on an alternative measure of openness suggest our baseline coefficients are quite robust, and may even understate the effect of openness on growth.

The last robustness check that we perform consists of testing our theory in a panel rather than a cross-sectional context. While the cross-sectional approach is preferable to capture the long term relationship between growth, openness and country size, and is now usual in the literature linking trade and productivity, a panel approach using decade averages may provide efficiency gains while still maintaining a relatively long horizon. We used the latest version of the PWT (version 6.1) and construct a panel of four decades spanning 1960-2000. We formulated a system of equations with two equations per period (one for the determination of growth and the other for trade openness), and constrained slope coefficients to equality across periods. We then ran 3SLS on this 8 equation system. ${ }^{34}$ Results are displayed in Tables 4 and 5 (columns 3 and 4). Once again, our results are qualitatively unchanged. In the growth equation (column 4 of Table 4), estimates on openness, log population and their interaction are very close to those obtained in the corresponding entry of Table 1 (column $4)$, and similarly for the effect of country size on openness in Table 5.

### 4.2 The Effects of Hypothetical Mergers

### 4.2.1 Effects on Expected 1960-1989 Growth

The parameter estimates presented in Tables 1 and 2 can be used to calculate, for pairs of adjacent countries, what their growth rate would have been had they formed a single country over the sample period under consideration. ${ }^{35}$ Namely, we can now calculate the impact of specific borders on growth, under alternative definitions of political integration. As described above, under a "size merger", which is reflected in equation (37), a political merger simply

[^18]entails full access to the neighbor's markets, without any change in the home country's $W_{a t}$ and $Z_{a t}$ variables. Under "full integration", reflected in equation (39), both hypothetically merged countries share the same $W_{a t}^{m}$ and $Z_{a t}^{m}$, and therefore the same growth rate under political integration. Since there is no a priori reason to prefer one definition over the other, we calculate the effect of borders under both definitions, and further decompose this effect into the direct positive effect of an increase in country size, the indirect negative effect via openness reduction, the ambiguous effect via the interaction term, and the steady-state determination effect.

Table 6 shows summary statistics for these various effects based on 123 hypothetical pairwise mergers. A salient feature of these statistics is the wide dispersion of the various effects. The pure size effect on growth, $\Delta G$, has a standard deviation of 0.377 and a positive mean of 0.123 percentage points of growth annually, suggesting that the average country would benefit from merging with a neighbor based on increased size alone. Indeed, the direct effect of size on growth, on average, more than outweighs the indirect effect via openness reduction (while the interaction effect is on average very close to zero). Under a full integration scenario, however, a typical country would lose slightly, on the order of $\Delta G^{m}=-0.112$ percentage points of annual growth. Since the difference between $\Delta G^{m}$ and $\Delta G$ is equal to the steadystate determination effect, the latter is on average negative (and equal to -0.235 ). Therefore, borders shield the average country from slow growing neighbors. There is, however, a wide dispersion of effects around this mean. This suggests that these simple summary statistics mask relevant country-specific features of the border effect on growth.

Figures 1 through 8 provide perhaps a more complete picture. They plot the distributions of the estimated effects. The total size effect $\Delta G$ is generally positive but moderate, in most cases smaller than 0.5 percentage points of annual growth. The effect of full political integration $\Delta G^{m}$ is more symmetrically distributed around zero, with slightly fatter tails. Turning to the decomposed effects confirms previous observations, namely that the interaction term effect is tightly distributed around zero, while the steady-state determination effect is slightly skewed, with a negative mean.

### 4.2.2 An Example: France and Italy

While these summary statistics and plotted distributions are useful, they are no substitute for the estimates obtained individually for each pair of adjacent countries. A close examination of these specific estimates reveals that their magnitudes are very sensible and that their signs are as expected. Small countries merging with large markets and poor countries merging with neighbors that exhibit superior steady-state determinants tend to gain. Large countries like the US tend to be indifferent to merging with small neighbors like Canada. The magnitudes of the gains and losses are commensurate with relative sizes and relative incomes.

To illustrate the results, we can examine more specifically the example of France and Italy (Table 7). The effect on France from merging with Italy would have been quite large and positive. We estimate that the total size effect would have resulted in a gain of 0.281 points of growth annually for France. To achieve a similar increase in growth via openness, France would have had to increase her trade to GDP ratio by 27.79 percentage points (for comparison, the average trade to GDP ratio of France over the sample period was $36 \%$ ). Since Italy started with a lower level of per capita income than France in 1960, but has a higher estimated steady-state income level given its observed steady-state determinants, France would also have gained from the steady-state determination effect. This effect alone would have accounted for $\Delta G^{m}-\Delta G=0.492$ additional points of growth. ${ }^{36}$

Turning to the effect on Italy from merging with France, it follows from what precedes that the steady-state determination effect would have been negative for Italy. Moreover, the positive size effect of a merger on Italian growth, equal to 0.237 , would not have been sufficient to outweigh the negative steady-state determination effect. Under full integration, Italy would have lost -0.316 points of growth annually. A possible interpretation of these results is that, if France and Italy could somehow have achieved the more restrictive form of political integration implied by the "size merger" definition, i.e. a removal of the border without changes in national savings rates, human capital, etc., both could have benefited in terms of growth.

[^19]Interested readers can ponder upon the estimated effects of their favorite hypothetical political merger among the 246 examples listed in Table 8.

### 4.2.3 Residual Effects

Section 3.3 above outlined a methodology to include the residuals from the growth and openness regressions into our analysis. Table 8 (columns 9 and 10) presents estimates of $\Delta G^{m e}$ and $\Delta G^{e}$ as in equations (40) and (41). The distribution of these effects is also displayed in Figures 9 and 10. Interestingly, the results do not change as much as expected given that the explained portions of growth and openness in the baseline regressions are only $60 \%$ and $50 \%$, respectively. The simple correlation of $\Delta G$ with and without the residual effect is 0.737 , while the corresponding figure for $\Delta G^{m}$ is 0.640 . Out of 246 mergers, accounting for the residual leads to a change in the sign of the effect in 31 cases for $\Delta G(12.6 \%$ of the cases) and 75 cases for $\Delta G^{m}(30.1 \%) .{ }^{37}$

Again, the case of France and Italy is illustrative (Table 7). Because France's explained annual growth falls short of its observed growth by 0.56 points, while Italy's observed and explained growth are about equal, accounting for the residual in the merger experiment is now slightly beneficial to Italy - which would have gained both under a size merger and full integration.

### 4.2.4 Effects on Steady-State Income Levels

Columns 11 and 12 of Table 8 presents, for each country pair, the estimated effect of a merger on the steady-state income level of country $a$, while the last row of Table 6 presents summary statistics for the steady-state level effects (the distribution of these level effects is displayed in Figures 11 and 12 for a size merger and full integration, respectively). On average, size mergers would raise a country's steady-state income level by 10.98 percentage points and full integration would reduce it by 2.07 percentage points. These averages reflect the generally positive effect of a size merger and the ambiguous effect of full integration. However, they again mask considerable case-specific differences. The effect of full integration ranges from

[^20]-421.07 percentage points (the effect on Malta from merging with Algeria - a small rich country merging with a relatively large poor country) to 325.63 percentage points (the effect on Papua New Guinea from merging with Australia - a small poor country merging with a rich country with five times its population). Logically, large effects such as these are found in cases where neighbors have very different sizes and income determinants.

More moderate effects are found in regions that are homogeneous in terms of income and size. For example, Table 7 shows that a size merger between France and Italy would have raised both countries' steady-state income levels by 25.1 percentage points for France and 21.12 points for Italy. Full integration would have reduced Italy's steady-state income by 15.89 percentage points. This partly reflects compounding the negative growth effect on Italy of full integration with France, as discussed earlier. The merger would raise France's steady-state income by 57.01 points, reflecting Italy's superior steady-state determinants.

### 4.3 Convergent Interests in Political Integration

An interesting application of our framework is to examine pairs of countries that would have both benefited from merging politically. As suggested above, it is much easier for two countries to have convergent interests in a size merger than in full integration, because the effect of the former is far more likely to be positive for any given country. Out of the 123 political mergers we considered in this paper, 94 entail growth gains for both country $a$ and country $b$ based on a size merger alone, and only 6 cases did the trade reduction effect dominate in both countries - so that both would have experienced reduced growth under a merger. These cases pertain to pairs of very small and already open countries, such as Singapore and Hong Kong or Jamaica and Haiti.

More interestingly perhaps, in only 14 cases would both countries in a merging pair have benefited from full integration in terms of economic growth. ${ }^{38}$ These pairs are listed in Table 9. Salient examples include Argentina and Chile, France and Germany, Canada and the US, India and Pakistan, as well as several country pairs involving Brazil. Of course, many more cases would entail a winner and a loser among the merging pair. 92 cases out of 123 entail exactly one country that would have gained from full political integration, while the

[^21]other would have lost, and in the remaining 17 cases both countries would have lost. The conclusion is that, in 109 of the 123 cases we considered, borders shield at least one country from the other.

An implication of these observations is that, when unions of country pairs are considered, it may be easier to gain mutual support for a form of political integration that shields countries from having to share their $Z_{a t}$ and $W_{a t}$ variables but focuses instead on taking advantage of scale effects, through the formation of free trade areas and the reduction of physical trading costs.

## 5 Conclusion

This paper provides a theoretical framework to understand the relationship between political borders and growth. We suggested that, whenever scale effects are present, political borders affect steady-state per capita income levels and transitional growth rates by reducing the extent of the market. We also pointed out that, in a world of more than two countries, the removal of only one border will result in trade reduction from the merging countries vis-àvis the rest of the world, with correspondingly adverse effects on growth and income. We examined formal conditions under which the extent of the market effect dominates the trade reduction effect, and discussed situations in which countries might differ in more that just size and openness levels.

We then derived an empirical specification directly from the theoretical model, and found strong empirical support for the predictions of our theory. Baseline parameter estimates from this empirical model were used to estimate, for specific countries, the growth effects of merging with another country. We have applied this framework to 123 pairs of adjacent countries and proximate islands. We found that full political integration with a neighbor would have a slight negative impact on the average country. This type of political integration, which entails full averaging of steady-state determinants across merging pairs, generally involves a winner and a loser - in only 14 of 123 cases would a merger have raised both country's growth rate. In contrast, countries would in general benefit from expanding the extent of their markets through deep economic integration with their neighbors, as shown by the prevalence of positive estimated effects under "size mergers". A limited form of integration that entails access to markets and a reduction of trade costs, to take advantage of scale effects, seems
more likely to benefit both countries in a pair than a form of integration that results in uniform growth determinants across country pairs.

Our framework can be extended in several directions. First, we have limited our investigation to hypothetical mergers involving only two countries. However, our framework is readily applicable to studying the growth effects of more than one political border. We could apply our methodology, for example, to the removal of all borders within Europe, in order to study the growth implications of proposals for European political integration. Our results for France and Germany suggest that both would have benefited, in terms of growth, from merging politically. Whether European countries would have benefited from the removal of all intra-European borders is an open and equally interesting question.

Second, our estimation method focuses exclusively on growth and income levels. There are obviously many other reasons, beyond growth, why countries would want to merge or stay separate. We can interpret our estimates of the growth effects of borders, whenever they are negative, as the amount of growth a country is willing to forego in order to avoid the non-economic costs of sharing a single polity with a neighbor. These may include increases in cultural, ethnic, religious or linguistic heterogeneity. Future work could relate changes in heterogeneity resulting from political integration to the magnitude of the growth costs or benefits. One interesting hypothesis to test is whether countries that remained separate despite potential growth effects of merging, have done so because political integration would have entailed large increases in heterogeneity.

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## Appendix 1

## Derivation of Equation (5)

At each time $t$, the resource constraint for each input $i \in[0, W]$ produced in a region $i$ belonging to a specific country $a$ of size $S_{a}$ is:

$$
\begin{equation*}
S_{a} x_{i a}(t)+\sum_{n \neq a} S_{n} x_{i n}(t)=K_{i}(t) \tag{50}
\end{equation*}
$$

Equations (4) and (50) imply that each region in country $a$ will use the same amount of domestically produced input $i$ :

$$
\begin{equation*}
x_{i a}(t)=\frac{K_{i}(t)}{S_{a}+\sum_{n \neq a} S_{n}\left(1-\xi_{a}-\xi_{n}\right)^{\frac{\alpha}{1-\alpha}}} \tag{51}
\end{equation*}
$$

On the other hand, each region of a country $b \neq a$ will use the following amount of input $i$ produced in country $a$ :

$$
\begin{equation*}
x_{i b}(t)=\frac{\left(1-\xi_{a}-\xi_{b}\right)^{\frac{\alpha}{1-\alpha}} K_{i}(t)}{S_{a}+\sum_{n \neq a} S_{n}\left(1-\xi_{a}-\xi_{n}\right)^{\frac{\alpha}{1-\alpha}}} \tag{52}
\end{equation*}
$$

By substituting (51) into (4) we obtain equation (5).

## Appendix 2

Reduced Form Parameters in equation (36):

$$
\left\{\begin{array}{c}
\gamma_{0}=\beta_{0}+\beta_{2} \alpha_{0} \\
\gamma_{1}=\beta_{1} \\
\gamma_{2}=\beta_{3}+\beta_{2} \alpha_{1}+\beta_{4} \alpha_{0} \\
\gamma_{3}=\beta_{4} \alpha_{1} \\
\gamma_{4}=\beta_{4} \alpha_{2} \\
\gamma_{5}=\beta_{4} \\
\gamma_{6}=\beta_{2} \alpha_{2} \\
\gamma_{7}=\beta_{5} \\
\mu_{i t}=\varepsilon_{i}+\beta_{2} \nu_{i t}
\end{array}\right.
$$

Table 1 - System Estimates of the Growth Equation

|  | 1960-1989 |  |  | 1960-1998 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 3SLS | 3SLS | SUR | 3SLS | 3SLS | SUR |
|  | (1) | (2) | (3) | (4) | (5) | (6) |
| Log population | $\begin{array}{r} 0.678 \\ (0.185) * * \end{array}$ | $\begin{array}{r} 1.337 \\ (\mathbf{0 . 2 5 4 ) * *} \\ \hline \end{array}$ | $\begin{array}{r} 0.263 \\ (0.134)^{* *} \end{array}$ | $\begin{array}{r} 0.472 \\ (0.249) * \end{array}$ | $\begin{array}{r} 1.387 \\ (0.311)^{* *} \end{array}$ | $\begin{array}{r} 0.130 \\ (0.136) \end{array}$ |
| Open* Log pop | $\begin{array}{r} -0.007 \\ (0.003) * * \end{array}$ | $\begin{array}{r} -0.007 \\ (0.004) * \end{array}$ | $\begin{array}{r} -0.003 \\ (0.002)^{*} \end{array}$ | $\begin{array}{r} -0.005 \\ (0.004) \end{array}$ | $\begin{array}{r} -0.009 \\ (0.006)^{*} \end{array}$ | $\begin{gathered} -\mathbf{0 . 0 0 1} \\ \mathbf{( 0 . 0 0 2}) \end{gathered}$ |
| Openness | $\begin{array}{r} 0.081 \\ (\mathbf{0 . 0 2 3})^{* *} \end{array}$ | $\begin{array}{r} 0.118 \\ (\mathbf{0 . 0 3 2})^{* *} \\ \hline \end{array}$ | $\begin{array}{r} 0.040 \\ (0.017)^{* *} \end{array}$ | $\begin{array}{r} 0.055 \\ (0.036)^{*} \end{array}$ | $\begin{array}{r} 0.124 \\ (0.048)^{* *} \end{array}$ | $\begin{array}{r} 0.010 \\ (0.019) \end{array}$ |
| Log 1960 per capita income | $\begin{array}{r} -1.120 \\ (0.269)^{* *} \end{array}$ | $\begin{array}{r} 0.120 \\ (0.205) \end{array}$ | $\begin{array}{r} -1.262 \\ (0.245)^{* *} \end{array}$ | $\begin{array}{r} -1.437 \\ (0.321) \\ \hline \end{array}$ | $\begin{array}{r} 0.322 \\ (0.216) \\ \hline \end{array}$ | $\begin{array}{r} -1.611 \\ (0.263) * * \end{array}$ |
| Fertility Rate | $\begin{array}{r} -0.185 \\ (0.121) \\ \hline \end{array}$ |  | $\begin{array}{r} -0.308 \\ (0.114)^{* *} \end{array}$ | $\begin{gathered} -0.601 \\ (0.152) \end{gathered}$ |  | $\begin{array}{r} -0.717 \\ (0.136)^{* *} \end{array}$ |
| Male human capital | $\begin{array}{r} 1.550 \\ (0.443)^{* *} \end{array}$ |  | $\begin{array}{r} 1.745 \\ (0.402)^{* *} \end{array}$ | $\begin{array}{r} 0.079 \\ (0.317) \end{array}$ |  | $\begin{array}{r} 0.010 \\ (0.295) \end{array}$ |
| Female human capital | $\begin{gathered} -1.183 \\ (0.472)^{* *} \end{gathered}$ |  | $\begin{array}{r} -1.415 \\ (0.433)^{* *} \end{array}$ | $\begin{array}{r} 0.162 \\ (0.395) \end{array}$ |  | $\begin{array}{r} 0.165 \\ (0.373) \\ \hline \end{array}$ |
| Government consumption ratio | $\begin{gathered} -0.053 \\ (0.020)^{* *} \end{gathered}$ |  | $\begin{array}{r} -0.061 \\ (0.019)^{* *} \end{array}$ | $\begin{gathered} -0.024 \\ (0.018) \end{gathered}$ |  | $\begin{array}{r} -0.031 \\ (0.017)^{*} \end{array}$ |
| Investment rate | $\begin{array}{r} 0.091 \\ (0.024)^{* *} \end{array}$ |  | $\begin{array}{r} 0.087 \\ (0.022)^{* *} \end{array}$ | $\begin{array}{r} 0.073 \\ (0.026) \end{array}$ |  | $\begin{array}{r} 0.075 \\ (0.021)^{* *} \end{array}$ |
| R-Squared | 0.558 | -0.277 | 0.683 | 0.662 | -0.221 | 0.726 |

(Standard errors in parentheses)
All regressions include an intercept term (output omitted).

* Denotes significance at the $90 \%$ confidence level; ** denotes significance at the $95 \%$ level.

92 observations in specifications for 1960-1989 and 77 observations in the specifications for 1960-1998.
Estimated jointly with the openness equation.

Table 2 - System Estimates of the Openness Equation

|  | $\mathbf{1 9 6 0 - 1 9 8 9}$ |  |  | $\mathbf{1 9 6 0 - 1 9 9 8}$ |  |  |
| :--- | ---: | :---: | ---: | ---: | ---: | ---: |
|  | 3SLS | 3SLS | SUR | 3SLS | 3SLS | SUR |
|  | $\mathbf{( 1 )}$ | $\mathbf{( 2 )}$ | $\mathbf{( 3 )}$ | $\mathbf{( 4 )}$ | $\mathbf{( 5 )}$ | $\mathbf{( 6 )}$ |
| Log population | $\mathbf{- 6 . 7 4 4}$ | $\mathbf{- 1 5 . 0 0 3}$ | $\mathbf{- 5 . 8 6 2}$ | $\mathbf{- 7 . 7 2 3}$ | $\mathbf{- 1 5 . 3 5 6}$ | $\mathbf{- 7 . 0 9 3}$ |
|  | $(\mathbf{( 2 . 6 7 1})^{* *}$ | $\mathbf{( 2 . 2 5 4 ) ^ { * * }}$ | $\mathbf{( 2 . 6 9 9 ) ^ { * * }}$ | $\mathbf{( 3 . 1 8 5 ) * *}$ | $\mathbf{( 2 . 9 3 5 ) ^ { * * }}$ | $\mathbf{( 3 . 2 1 3 ) * *}$ |
| Log 1960 per | 1.335 | 3.378 | 1.888 | 1.526 | 0.753 | 1.696 |
| capita income | $(3.868)$ | $(4.139)$ | $(3.902)$ | $(4.802)$ | $(5.314)$ | $(4.826)$ |
| Log land | -9.868 |  | -10.537 | -10.511 |  | -11.271 |
| area | $(2.124)^{* *}$ |  | $(2.179)^{* *}$ | $(2.542)^{* *}$ |  | $(2.596)^{* *}$ |
| Terms of | -45.202 |  | 48.984 | 373.600 |  | 377.467 |
| trade shocks | $(205.930)$ |  | $(221.254)$ | $(291.622)$ |  | $(302.285)$ |
| Oil | 13.999 |  | 9.771 | -13.199 |  | -15.031 |
| dummy | $(21.898)$ |  | $(23.596)$ | $(28.132)$ |  | $(29.393)$ |
| Landlock | -2.472 |  | 1.807 | -6.386 |  | -5.702 |
| Dummy | $(8.889)$ |  | $(9.602)$ | $(10.285)$ |  | $(10.772)$ |
| Island | 3.186 |  | 4.337 | 12.643 |  | 11.276 |
| Dummy | $(7.766)$ |  | $(8.351)$ | $(9.934)$ |  | $(10.385)$ |
| R-Squared | 0.508 | 0.333 | 0.511 | 0.506 | 0.270 | 0.507 |

(Standard errors in parentheses)
All regressions include an intercept term (output omitted).

* Denotes significance at the $90 \%$ confidence level; ${ }^{* *}$ denotes significance at the $95 \%$ level.

92 observations in all specifications for 1960-1989 and 77 observations for 1960-1998. Estimated jointly with the growth equation.

Table 3 - First stage F-Tests for the Instruments

| Specification | Test | Openness | Openness*Log <br> Population |
| :--- | :--- | ---: | ---: |
| $(1)$ | $\mathrm{F}(11,73)$ | 3.11 | 3.09 |
|  | p -value) | $(0.002)$ | $(0.002)$ |
| $(2)$ | $\mathrm{F}(11,78)$ | 3.55 | 3.72 |
|  | p -value $)$ | $(0.0005)$ | $(0.0003)$ |

Table 4 - Growth Equation - Robustness Checks (estimator: 3SLS)

|  | Alcala-Ciccone Openness Measure |  | 4 Decade Panel (1960-2000) |  |
| :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) |
| Log Population | $\begin{array}{r} 1.174 \\ (0.293)^{* *} \end{array}$ | $\begin{array}{r} 0.447 \\ (0.205) * * \\ \hline \end{array}$ | $\begin{array}{r} 0.761 \\ (0.144) * * \end{array}$ | $\begin{array}{r} 0.437 \\ \left(\mathbf{0 . 1 4 4 )}{ }^{* *}\right. \end{array}$ |
| Open*Log Pop | $\begin{array}{r} -0.019 \\ (0.009)^{* *} \end{array}$ | $\begin{array}{r} -0.012 \\ (0.006)^{*} \\ \hline \end{array}$ | $\begin{array}{r} -0.009 \\ (0.002) * * \end{array}$ | $\begin{array}{r} -0.005 \\ (0.002)^{* *} \\ \hline \end{array}$ |
| Openness | $\begin{array}{r} 0.233 \\ (0.074)^{* *} \\ \hline \end{array}$ | $\begin{array}{r} 0.122 \\ (0.050) * * \\ \hline \end{array}$ | $\begin{array}{r} 0.096 \\ (0.019) * * \end{array}$ | $\begin{array}{r} 0.055 \\ (0.019) * * \\ \hline \end{array}$ |
| Log Per Capita Income 1960 | $\begin{array}{r} -0.489 \\ (0.267)^{*} \\ \hline \end{array}$ | $\begin{array}{r} -1.514 \\ (0.269)^{* *} \end{array}$ | $\begin{array}{r} 0.167 \\ (0.112) \end{array}$ | $\begin{array}{r} -1.056 \\ (0.227)^{* *} \\ \hline \end{array}$ |
| Fertility Rate |  | $\begin{array}{r} -0.343 \\ (0.128)^{* *} \\ \hline \end{array}$ |  | $\begin{array}{r} -0.408 \\ (0.109)^{* *} \end{array}$ |
| Male Human Capital |  | $\begin{array}{r} 1.619 \\ (0.467)^{* *} \end{array}$ |  | $\begin{array}{r} 0.341 \\ (0.249) \\ \hline \end{array}$ |
| Female Human Capital |  | $\begin{array}{r} -1.229 \\ (0.489)^{* *} \end{array}$ |  | $\begin{array}{r} -0.232 \\ (0.294) \\ \hline \end{array}$ |
| Government Consumption Ratio |  | $\begin{array}{r} -0.060 \\ (0.021)^{* *} \end{array}$ |  | $\begin{array}{r} -0.032 \\ (0.012)^{* *} \end{array}$ |
| Investment Rate |  | $\begin{array}{r} 0.083 \\ (0.027)^{* *} \\ \hline \end{array}$ |  | $\begin{array}{r} 0.094 \\ (0.016)^{* *} \end{array}$ |
| \# countries (\# periods) | 88 (1) | 88 (1) | 99 (4) | 79 (4) |
| R-squared | -0.390 | 0.596 | $\begin{gathered} \hline-0.140 .01 \\ -0.18-0.08 \end{gathered}$ | $\begin{aligned} & \hline 0.390 .21 \\ & 0.460 .21 \end{aligned}$ |

Standard errors in parentheses
Regression in column 4 includes period fixed effects (output omitted).
Other regressions include an intercept term (output omitted).

* Denotes significance at the $90 \%$ confidence level; $* *$ denotes significance at the $95 \%$ level.

Table 5-Openness Equation - Robustness Checks (estimator: 3SLS)

|  | Alcala-Ciccone Openness Measure |  | 4 Decade Panel (1960-2000) |  |
| :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) |
| Log Population | $\begin{array}{r} -7.692 \\ (1.852)^{*} * \end{array}$ | $\begin{array}{r} \hline-1.895 \\ (2.293) \\ \hline \end{array}$ | $\begin{array}{r} -10.364 \\ (1.218) * * \end{array}$ | $\begin{array}{r} -8.541 \\ (1.741)^{*} * \end{array}$ |
| Log Per Capita Income 1960 | $\begin{array}{r} 11.608 \\ (3.275)^{* *} \\ \hline \end{array}$ | $\begin{array}{r} 10.761 \\ (3.288)^{* *} \\ \hline \end{array}$ | $\begin{array}{r} 8.570 \\ (1.742)^{* *} \\ \hline \end{array}$ | $\begin{array}{r} 2.781 \\ (2.038) \\ \hline \end{array}$ |
| Log Land Area |  | $\begin{array}{r} -7.302 \\ (1.854)^{* *} \end{array}$ |  | $\begin{array}{r} -5.611 \\ (1.454)^{* *} \end{array}$ |
| Terms of Trade Shocks |  | $\begin{array}{r} 17.399 \\ (179.445) \\ \hline \end{array}$ |  | $\begin{array}{r} 38.631 \\ (12.662)^{* *} \end{array}$ |
| Oil Dummy |  | $\begin{array}{r} 6.263 \\ (19.024) \\ \hline \end{array}$ |  | $\begin{array}{r} 11.128 \\ (9.927) \\ \hline \end{array}$ |
| Landlock Dummy |  | $\begin{array}{r} \hline-0.138 \\ (7.921) \\ \hline \end{array}$ |  | $\begin{array}{r} -3.869 \\ (6.130) \\ \hline \end{array}$ |
| Island Dummy |  | $\begin{array}{r} -0.776 \\ (6.830) \\ \hline \end{array}$ |  | $\begin{array}{r} 0.132 \\ (5.622) \\ \hline \end{array}$ |
| \# countries (\# periods) | 88 (1) | 88 (1) | 99 (4) | 79 (4) |
| R-squared | 0.239 | 0.367 | $\begin{aligned} & \hline 0.270 .45 \\ & 0.350 .08 \end{aligned}$ | $\begin{array}{ll} \hline 0.53 & 0.57 \\ 0.50 & 0.37 \end{array}$ |

Standard errors in parentheses
Regression in column 4 includes period fixed effect (output omitted).
Other regressions include an intercept term (output omitted).

* Denotes significance at the $90 \%$ confidence level; $* *$ denotes significance at the $95 \%$ level.

Table 6 - Summary Statistics of the Effects of Border Removals
Based on Tables 1 and 2, column (1) estimates

| Variable | Mean | Std. Dev. | Min | Max |
| :--- | ---: | ---: | ---: | ---: |
| Observed average growth | 2.127 | 1.671 | -1.231 | 6.580 |
| Fitted growth | 2.167 | 1.369 | -0.006 | 6.150 |
| Direct effect of size | 0.745 | 0.743 | 0.005 | 3.452 |
| Indirect effect via openness reduction | -0.601 | 0.600 | -2.784 | -0.004 |
| Effect via change in interaction term | -0.021 | 0.405 | -2.400 | 1.056 |
| Steady-state determination effect | -0.235 | 0.678 | -3.442 | 2.350 |
| $\Delta \mathrm{G}$ ("size merger") | 0.123 | 0.377 | -1.733 | 1.367 |
| $\Delta \mathrm{G}^{\mathrm{M}}$ ("full integration") | -0.112 | 0.914 | -4.965 | 3.350 |
| Openness equivalent ("size merger") | 10.184 | 22.517 | -40.281 | 184.332 |
| Openness equivalent ("full integration") | 0.403 | 43.089 | -144.379 | 315.772 |
| $\Delta \mathrm{G}^{\mathrm{e}}$ ("size merger" with residual effect) | 0.097 | 0.524 | -4.214 | 1.511 |
| $\Delta \mathrm{G}^{\text {me }}$ ("full integration" with residual effect) | -0.092 | 1.047 | -4.036 | 4.220 |
| $\Delta \mathrm{SSY}^{\text {(steady-state level effect of a "size merger") }(\%)}$ | 10.976 | 33.623 | -154.651 | 121.956 |
| $\Delta \mathrm{SSY}^{\mathrm{m}}$ (steady-state level effect of "full integration") $(\%)$ | -2.068 | 83.400 | -421.068 | 325.630 |

(Based on 246 effects calculated from 123 hypothetical political mergers)

Table 7-An Example: France and Italy

| Effect on (country a): | France | Italy |
| :--- | ---: | ---: |
| of merging with (country b): | Italy | France |
| Observed Growth (country a) | 2.936 | 3.404 |
| Fitted Growth (country a) | 2.374 | 3.464 |
| Direct effect of size | 0.491 | 0.451 |
| Indirect effect via openness | -0.396 | -0.364 |
| Effect via change in interaction term | 0.186 | 0.149 |
| Steady-state determination effect | 0.492 | -0.553 |
| $\Delta \mathrm{G}$ ("size merger") | 0.281 | 0.237 |
| $\Delta \mathrm{G}^{\mathrm{m}}$ ("full integration") | 0.773 | -0.316 |
| Openness equivalent ("size merger") | 27.789 | 24.300 |
| Openness equivalent ("full integration") | 76.423 | -32.492 |
| $\Delta \mathrm{G}^{\mathrm{e}}$ ("size merger" with residual effect) | 0.294 | 0.265 |
| $\Delta \mathrm{G}^{\mathrm{me}}$ ("full integration" with residual effect) | 0.474 | 0.006 |
| $\Delta \mathrm{SSY}$ (steady-state level effect of a "size merger") (\%) | 25.099 | 21.122 |
| $\Delta \mathrm{SSY}$ (steady-state level effect of "full integration") $(\%)$ | 57.011 | -15.894 |

Table 8 - Country-Specific Merger Estimates

| Country a | Country b | a's Fitted Growth | $\Delta \mathbf{G}$ | $\Delta \mathbf{G}^{\mathbf{m}}$ | direct | indirect | interact | $\Delta G^{\text {e }}$ | $\Delta \mathbf{G}^{\text {me }}$ | $\begin{gathered} \Delta \mathbf{S S Y} \\ (\%) \\ \hline \end{gathered}$ | $\begin{gathered} \Delta \mathbf{S S Y}^{\mathrm{m}} \\ (\%) \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 |
| Algeria | Mali | 1.124 | 0.103 | -0.178 | 0.232 | -0.187 | 0.058 | 0.076 | -0.392 | 9.209 | -37.268 |
|  | Malta | 1.124 | 0.015 | 0.061 | 0.035 | -0.028 | 0.008 | 0.011 | 0.026 | 1.330 | 4.839 |
|  | Niger | 1.124 | 0.087 | -0.170 | 0.198 | -0.160 | 0.049 | 0.064 | -0.418 | 7.789 | -32.517 |
|  | Tunisia | 1.124 | 0.101 | 0.219 | 0.228 | -0.184 | 0.057 | 0.074 | 0.417 | 9.015 | 8.828 |
| Argentina | Bolivia | 0.987 | 0.075 | 0.134 | 0.122 | -0.099 | 0.051 | 0.089 | 0.136 | 6.658 | 0.786 |
|  | Brazil | 0.987 | 0.779 | 1.340 | 1.107 | -0.893 | 0.565 | 0.907 | 2.301 | 69.529 | 56.585 |
|  | Chile | 0.987 | 0.143 | 0.110 | 0.230 | -0.186 | 0.098 | 0.169 | 0.452 | 12.734 | -0.268 |
|  | Paraguay | 0.987 | 0.044 | 0.011 | 0.073 | -0.059 | 0.030 | 0.053 | 0.157 | 3.943 | -5.183 |
|  | Uruguay | 0.987 | 0.047 | -0.008 | 0.077 | -0.062 | 0.032 | 0.056 | 0.110 | 4.167 | -1.919 |
| Australia | Fiji | 2.291 | 0.021 | -0.059 | 0.035 | -0.028 | 0.015 | 0.018 | -0.072 | 1.915 | -7.981 |
|  | Indonesia | 2.291 | 1.232 | 0.531 | 1.613 | -1.301 | 0.920 | 1.065 | 0.781 | 109.960 | -128.110 |
|  | New Zealand | 2.291 | 0.089 | -0.030 | 0.142 | -0.115 | 0.061 | 0.074 | -0.156 | 7.905 | -2.256 |
|  | Papua New Guinea | 2.291 | 0.082 | -0.550 | 0.131 | -0.106 | 0.056 | 0.068 | -0.449 | 7.279 | -63.385 |
|  | Sri Lanka | 2.291 | 0.310 | -0.587 | 0.473 | -0.382 | 0.218 | 0.261 | -0.882 | 27.626 | -105.326 |
| Austria | rmany, Fed. Rep. | 4.128 | 0.403 | -1.464 | 1.496 | -1.207 | 0.114 | 0.460 | -0.345 | 35.996 | -108.690 |
|  | Italy | 4.128 | 0.379 | -0.535 | 1.436 | -1.159 | 0.101 | 0.433 | 0.405 | 33.817 | -58.122 |
|  | Switzerland | 4.128 | 0.067 | -0.404 | 0.408 | -0.329 | -0.012 | 0.083 | -0.588 | 6.016 | -5.444 |
| Bangladesh | India | 0.851 | 0.926 | 2.501 | 1.509 | -1.217 | 0.634 | 1.273 | 0.579 | 82.667 | 204.027 |
| Barbados | Colombia | 4.310 | -1.673 | -2.428 | 3.081 | -2.485 | -2.268 | -1.041 | -1.337 | -149.268 | -261.782 |
|  | Trinidad \& Tobago | 4.310 | -0.806 | -2.249 | 1.100 | -0.887 | -1.019 | -0.581 | -2.359 | -71.976 | -138.866 |
|  | Venezuela | 4.310 | -1.549 | -3.881 | 2.653 | -2.140 | -2.063 | -1.006 | -3.836 | -138.287 | -261.606 |
| Belgium | France | 3.051 | 0.199 | -0.530 | 1.250 | -1.008 | -0.043 | -0.217 | -0.008 | 17.731 | -42.409 |
|  | Germany, Fed. Rep. | 3.051 | 0.225 | -0.493 | 1.340 | -1.081 | -0.034 | -0.221 | -0.323 | 20.059 | -28.483 |
|  | Netherlands | 3.051 | 0.056 | -0.025 | 0.585 | -0.472 | -0.058 | -0.139 | -0.407 | 4.971 | 3.541 |
| Benin | Niger | -0.006 | 0.063 | -0.012 | 0.661 | -0.533 | -0.064 | 0.190 | 0.160 | 5.653 | -39.016 |
|  | Togo | -0.006 | 0.028 | 0.810 | 0.398 | -0.321 | -0.049 | 0.104 | 0.784 | 2.506 | 39.002 |
| Bolivia | Argentina | 1.951 | 0.578 | -0.830 | 1.266 | -1.021 | 0.333 | 0.522 | -0.819 | 51.610 | 50.535 |
|  | Brazil | 1.951 | 1.152 | 0.589 | 2.132 | -1.720 | 0.739 | 1.057 | 1.719 | 102.778 | 95.070 |


| Country a | Country b | a's Fitted Growth | $\Delta \mathbf{G}$ | $\Delta \mathbf{G}^{\mathbf{m}}$ | direct | indirect | interact | $\Delta G^{\text {e }}$ | $\Delta \mathbf{G}^{\text {me }}$ | $\begin{gathered} \Delta \text { SSY } \\ (\%) \end{gathered}$ | $\begin{gathered} \Delta \mathbf{S S Y}^{\mathbf{m}} \\ (\%) \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 |
| Bolivia | Chile | 1.951 | 0.320 | -0.704 | 0.780 | -0.629 | 0.169 | 0.285 | 0.042 | 28.560 | 8.816 |
|  | Paraguay | 1.951 | 0.115 | -0.230 | 0.314 | -0.254 | 0.054 | 0.101 | 0.224 | 10.250 | -19.665 |
|  | Peru | 1.951 | 0.411 | -0.162 | 0.962 | -0.776 | 0.225 | 0.369 | -0.640 | 36.723 | 30.274 |
| Botswana | South Africa | 1.341 | 0.666 | -0.241 | 2.438 | -1.967 | 0.195 | -0.191 | -3.395 | 59.460 | 117.516 |
|  | Zimbabwe | 1.341 | 0.278 | -0.182 | 1.509 | -1.217 | -0.014 | -0.253 | -4.036 | 24.780 | 39.610 |
| Brazil | Argentina | 2.488 | 0.149 | -0.161 | 0.164 | -0.132 | 0.118 | 0.132 | -0.382 | 13.330 | 14.310 |
|  | Bolivia | 2.488 | 0.041 | 0.052 | 0.045 | -0.037 | 0.032 | 0.036 | -0.009 | 3.658 | 3.047 |
|  | Colombia | 2.488 | 0.136 | 0.114 | 0.149 | -0.120 | 0.107 | 0.120 | 0.097 | 12.144 | 9.214 |
|  | Guyana | 2.488 | 0.018 | 0.012 | 0.019 | -0.016 | 0.014 | 0.015 | -0.021 | 1.562 | 0.960 |
|  | Paraguay | 2.488 | 0.029 | 0.017 | 0.032 | -0.026 | 0.023 | 0.025 | 0.025 | 2.577 | 0.640 |
|  | Peru | 2.488 | 0.094 | 0.070 | 0.104 | -0.084 | 0.074 | 0.083 | -0.161 | 8.404 | 7.790 |
|  | Uruguay | 2.488 | 0.030 | -0.029 | 0.033 | -0.027 | 0.023 | 0.026 | -0.027 | 2.664 | 1.431 |
|  | Venezuela | 2.488 | 0.081 | -0.115 | 0.089 | -0.072 | 0.063 | 0.071 | -0.232 | 7.202 | 10.716 |
| Cameroon | Central Afr. Rep. | 1.371 | 0.056 | -0.317 | 0.183 | -0.148 | 0.020 | 0.061 | -0.769 | 4.957 | -26.023 |
|  | Congo | 1.371 | 0.040 | -0.230 | 0.132 | -0.107 | 0.014 | 0.044 | -0.162 | 3.528 | -9.715 |
| Canada | U.S.A | 1.777 | 1.367 | 1.288 | 1.608 | -1.297 | 1.056 | 0.888 | -0.738 | 121.956 | 143.482 |
| Central Afr. Rep. | Cameroon | 0.137 | 0.289 | 0.918 | 1.042 | -0.841 | 0.087 | 0.194 | 2.466 | 25.775 | 74.775 |
|  | Congo | 0.137 | 0.075 | 0.037 | 0.357 | -0.288 | 0.006 | 0.043 | 1.064 | 6.736 | 23.269 |
|  | Zaire | 0.137 | 0.578 | 0.650 | 1.699 | -1.370 | 0.249 | 0.423 | 0.151 | 51.575 | 25.503 |
| Chile | Argentina | 0.991 | 0.377 | 0.107 | 0.866 | -0.699 | 0.209 | 0.427 | -0.646 | 33.621 | 43.039 |
|  | Bolivia | 0.991 | 0.103 | 0.257 | 0.273 | -0.220 | 0.050 | 0.119 | -0.101 | 9.193 | 2.374 |
|  | Peru | 0.991 | 0.253 | 0.481 | 0.617 | -0.497 | 0.134 | 0.289 | -0.525 | 22.615 | 24.426 |
| Colombia | Barbados | 1.880 | 0.011 | 0.002 | 0.022 | -0.018 | 0.007 | 0.013 | -0.001 | 1.021 | 1.057 |
|  | Brazil | 1.880 | 0.740 | 0.723 | 1.178 | -0.950 | 0.512 | 0.811 | 0.802 | 66.018 | 69.280 |
|  | Ecuador | 1.880 | 0.102 | 0.131 | 0.192 | -0.155 | 0.065 | 0.114 | 0.096 | 9.129 | 8.639 |
|  | Panama | 1.880 | 0.033 | -0.003 | 0.063 | -0.051 | 0.021 | 0.036 | -0.016 | 2.913 | -0.694 |
|  | Peru | 1.880 | 0.192 | 0.077 | 0.350 | -0.282 | 0.124 | 0.213 | -0.576 | 17.120 | 14.304 |
|  | Venezuela | 1.880 | 0.164 | -0.466 | 0.302 | -0.244 | 0.106 | 0.183 | -0.803 | 14.655 | 21.271 |
| Congo | Cameroon | 0.455 | 0.190 | 0.687 | 1.274 | -1.027 | -0.056 | -0.223 | -0.063 | 16.991 | 15.983 |
|  | Central Afr. Rep. | 0.455 | 0.057 | -0.282 | 0.640 | -0.516 | -0.067 | -0.151 | -2.072 | 5.053 | -51.830 |
|  | Zaire | 0.455 | 0.423 | 0.330 | 1.962 | -1.583 | 0.044 | -0.213 | -2.862 | 37.755 | -46.654 |
| Costa Rica | Nicaragua | 1.272 | -0.033 | -0.561 | 0.562 | -0.454 | -0.142 | 0.065 | -1.446 | -2.981 | -64.023 |


| Country a | Country b | a's Fitted Growth | $\Delta \mathbf{G}$ | $\Delta \mathbf{G}^{\mathbf{m}}$ | direct | indirect | interact | $\Delta \mathbf{G}^{\text {e }}$ | $\Delta \mathrm{G}^{\mathrm{me}}$ | $\begin{gathered} \Delta \text { SSY } \\ (\%) \\ \hline \end{gathered}$ | $\begin{gathered} \Delta \mathbf{S S Y}^{\mathrm{m}} \\ (\%) \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 |
| Costa Rica | Panama | 1.272 | -0.031 | -0.001 | 0.445 | -0.359 | -0.118 | 0.046 | -0.077 | -2.807 | -12.696 |
| Cyprus | Greece | 4.934 | -0.524 | -1.323 | 1.870 | -1.509 | -0.886 | -0.217 | -0.647 | -46.761 | -115.560 |
|  | Israel | 4.934 | -0.423 | -2.252 | 1.241 | -1.001 | -0.663 | -0.219 | -1.346 | -37.718 | -156.765 |
|  | Turkey | 4.934 | -0.533 | -2.165 | 2.812 | -2.268 | -1.077 | -0.073 | -2.068 | -47.599 | -215.453 |
| Denmark | Germany, Fed. Rep. | 2.198 | 0.261 | 0.236 | 1.747 | -1.409 | -0.077 | 0.543 | 0.171 | 23.294 | 18.404 |
|  | Iceland | 2.198 | 0.000 | -0.267 | 0.029 | -0.023 | -0.006 | 0.004 | -0.255 | -0.040 | -24.818 |
| Dominican Rep. | Haiti | 2.074 | 0.018 | -0.538 | 0.486 | -0.392 | -0.076 | 0.159 | -1.368 | 1.615 | -61.046 |
|  | Jamaica | 2.074 | 0.004 | -0.048 | 0.247 | -0.199 | -0.044 | 0.075 | -0.495 | 0.314 | 10.464 |
| Ecuador | Colombia | 2.127 | 0.310 | -0.117 | 1.017 | -0.820 | 0.113 | 0.444 | 0.173 | 27.662 | 0.758 |
|  | Peru | 2.127 | 0.226 | -0.273 | 0.796 | -0.642 | 0.072 | 0.331 | -1.188 | 20.148 | -1.150 |
| El Salvador | Guatemala | 0.841 | -0.024 | -0.168 | 0.646 | -0.521 | -0.149 | 0.159 | 0.000 | -2.156 | -5.590 |
|  | Honduras | 0.841 | -0.025 | -0.055 | 0.413 | -0.333 | -0.105 | 0.093 | -0.056 | -2.204 | -17.273 |
| Fiji | Australia | 2.621 | -0.430 | -0.388 | 2.191 | -1.767 | -0.853 | -0.171 | 0.108 | -38.344 | 93.211 |
|  | New Zealand | 2.621 | -0.359 | -0.653 | 1.256 | -1.013 | -0.602 | -0.211 | -0.848 | -32.053 | 63.566 |
| Finland | Norway | 2.486 | 0.090 | 0.201 | 0.413 | -0.333 | 0.010 | 0.118 | -0.140 | 8.018 | 20.619 |
|  | Sweden | 2.486 | 0.165 | -0.414 | 0.678 | -0.547 | 0.034 | 0.212 | -0.725 | 14.727 | -12.821 |
| France | Belgium | 2.374 | 0.063 | 0.147 | 0.118 | -0.095 | 0.041 | 0.067 | -0.026 | 5.663 | 12.174 |
|  | Germany, Fed. Rep. | 2.374 | 0.303 | 0.245 | 0.526 | -0.425 | 0.202 | 0.317 | 0.020 | 27.083 | 28.662 |
|  | Italy | 2.374 | 0.281 | 0.773 | 0.491 | -0.396 | 0.186 | 0.294 | 0.474 | 25.099 | 57.011 |
|  | Spain | 2.374 | 0.197 | 0.470 | 0.352 | -0.284 | 0.129 | 0.206 | 0.509 | 17.566 | 21.476 |
|  | Switzerland | 2.374 | 0.042 | 0.142 | 0.078 | -0.063 | 0.026 | 0.044 | -0.072 | 3.706 | 18.897 |
|  | United Kingdom | 2.374 | 0.286 | -0.364 | 0.498 | -0.402 | 0.189 | 0.299 | -0.140 | 25.496 | -23.711 |
| Germany, Fed. Rep. | Austria | 2.432 | 0.039 | 0.231 | 0.079 | -0.064 | 0.023 | 0.038 | 0.099 | 3.438 | 18.175 |
|  | Belgium | 2.432 | 0.049 | 0.125 | 0.101 | -0.082 | 0.030 | 0.049 | 0.026 | 4.409 | 8.848 |
|  | Denmark | 2.432 | 0.026 | 0.001 | 0.054 | -0.044 | 0.016 | 0.026 | 0.029 | 2.329 | 0.341 |
|  | France | 2.432 | 0.217 | 0.187 | 0.419 | -0.338 | 0.136 | 0.213 | 0.387 | 19.352 | 11.411 |
|  | Netherlands | 2.432 | 0.066 | 0.180 | 0.135 | -0.109 | 0.040 | 0.065 | 0.010 | 5.924 | 14.802 |
|  | Switzerland | 2.432 | 0.032 | 0.125 | 0.066 | -0.054 | 0.019 | 0.032 | -0.022 | 2.865 | 14.891 |
| Ghana | Togo | 1.545 | 0.044 | 0.091 | 0.157 | -0.126 | 0.013 | 0.076 | 0.308 | 3.899 | -3.164 |
| Greece | Cyprus | 3.605 | 0.009 | 0.006 | 0.046 | -0.037 | 0.000 | 0.019 | -0.009 | 0.760 | 0.345 |
|  | Turkey | 3.605 | 0.324 | -0.634 | 1.121 | -0.904 | 0.108 | 0.577 | -1.059 | 28.947 | -75.536 |
|  | Yugoslavia | 3.605 | 0.211 | 0.321 | 0.811 | -0.654 | 0.054 | 0.393 | -0.416 | 18.797 | 22.824 |


| Country a | Country b | a's Fitted Growth | $\Delta \mathrm{G}$ | $\Delta \mathbf{G}^{\mathbf{m}}$ | direct | indirect | interact | $\Delta \mathbf{G}^{\text {e }}$ | $\Delta \mathbf{G}^{\text {me }}$ | $\begin{gathered} \Delta \text { SSY } \\ (\%) \end{gathered}$ | $\begin{gathered} \hline \Delta \mathbf{S S Y}^{\mathrm{m}} \\ (\%) \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 |
| Guatemala | El Salvador | 0.815 | 0.053 | -0.142 | 0.360 | -0.290 | -0.017 | 0.155 | -0.029 | 4.692 | -18.431 |
|  | Honduras | 0.815 | 0.043 | 0.046 | 0.304 | -0.245 | -0.016 | 0.129 | 0.039 | 3.813 | -9.188 |
|  | Mexico | 0.815 | 0.436 | 1.022 | 1.629 | -1.314 | 0.121 | 0.900 | 1.416 | 38.943 | 140.856 |
| Guinea Bissau | Senegal | 2.243 | -0.141 | -2.056 | 1.444 | -1.165 | -0.420 | 0.472 | -0.780 | -12.559 | -117.343 |
| Guyana | Brazil | 1.286 | 0.726 | 1.214 | 3.410 | -2.750 | 0.067 | -0.846 | 4.220 | 64.789 | 119.398 |
|  | Venezuela | 1.286 | 0.150 | -0.882 | 1.981 | -1.598 | -0.233 | -0.764 | 0.851 | 13.365 | 53.913 |
| Haiti | Dominican Rep. | 1.288 | -0.008 | 0.247 | 0.477 | -0.385 | -0.100 | 0.207 | 1.433 | -0.727 | 34.801 |
|  | Jamaica | 1.288 | -0.010 | 0.156 | 0.237 | -0.191 | -0.055 | 0.097 | 0.315 | -0.849 | 37.999 |
| Honduras | El Salvador | 1.314 | 0.049 | -0.527 | 0.576 | -0.465 | -0.062 | 0.108 | -0.320 | 4.396 | -27.723 |
|  | Guatemala | 1.314 | 0.077 | -0.452 | 0.753 | -0.607 | -0.068 | 0.153 | -0.197 | 6.879 | -6.797 |
|  | Nicaragua | 1.314 | 0.029 | -0.472 | 0.414 | -0.334 | -0.051 | 0.071 | -0.891 | 2.580 | -20.166 |
| Hong Kong | Singapore | 5.622 | -0.098 | -0.251 | 0.282 | -0.228 | -0.153 | -0.288 | -0.878 | -8.755 | -32.005 |
| Iceland | Denmark | 3.396 | -0.363 | -1.465 | 2.170 | -1.750 | -0.783 | 0.002 | -1.270 | -32.409 | -100.837 |
|  | Norway | 3.396 | -0.367 | -0.441 | 2.020 | -1.629 | -0.758 | -0.027 | -0.227 | -32.768 | -27.661 |
|  | Sweden | 3.396 | -0.340 | -1.553 | 2.489 | -2.008 | -0.822 | 0.079 | -1.189 | -30.363 | -96.927 |
| India | Bangladesh | 3.424 | 0.097 | -0.072 | 0.092 | -0.074 | 0.079 | 0.087 | 0.199 | 8.652 | -3.850 |
|  | Pakistan | 3.424 | 0.092 | 0.157 | 0.087 | -0.070 | 0.075 | 0.082 | 0.329 | 8.170 | 12.410 |
|  | Sri Lanka | 3.424 | 0.029 | 0.028 | 0.027 | -0.022 | 0.023 | 0.026 | 0.050 | 2.561 | 3.923 |
| Indonesia | Australia | 3.392 | 0.052 | -0.569 | 0.079 | -0.064 | 0.037 | 0.048 | -0.616 | 4.658 | 23.824 |
|  | Malaysia | 3.392 | 0.048 | 0.051 | 0.072 | -0.058 | 0.034 | 0.044 | 0.055 | 4.259 | 13.995 |
|  | Papua New Guinea | 3.392 | 0.017 | -0.046 | 0.026 | -0.021 | 0.012 | 0.016 | -0.045 | 1.533 | -2.267 |
|  | Philippines | 3.392 | 0.135 | -0.069 | 0.201 | -0.162 | 0.096 | 0.125 | -0.329 | 12.075 | 10.221 |
|  | Sri Lanka | 3.392 | 0.052 | -0.049 | 0.078 | -0.063 | 0.037 | 0.048 | -0.129 | 4.614 | 4.570 |
| Iran | Pakistan | 1.219 | 0.448 | 1.015 | 0.801 | -0.646 | 0.293 | 0.467 | 1.632 | 39.960 | 12.156 |
|  | Turkey | 1.219 | 0.300 | 1.072 | 0.559 | -0.451 | 0.192 | 0.313 | 1.509 | 26.740 | 65.722 |
| Ireland | United Kingdom | 2.333 | 0.272 | -0.982 | 1.986 | -1.602 | -0.112 | -0.010 | -0.950 | 24.301 | -18.009 |
| Israel | Cyprus | 2.515 | -0.017 | 0.168 | 0.137 | -0.111 | -0.044 | 0.007 | 0.058 | -1.539 | 5.720 |
|  | Jordan | 2.515 | -0.037 | -0.721 | 0.360 | -0.291 | -0.107 | 0.027 | -0.357 | -3.347 | -99.526 |
|  | Syria | 2.515 | -0.049 | -0.865 | 0.811 | -0.654 | -0.206 | 0.096 | -0.350 | -4.398 | -124.013 |
| Italy | Austria | 3.464 | 0.043 | 0.129 | 0.087 | -0.071 | 0.026 | 0.048 | 0.014 | 3.820 | 13.090 |
|  | France | 3.464 | 0.237 | -0.316 | 0.451 | -0.364 | 0.149 | 0.265 | 0.006 | 21.122 | -15.894 |
|  | Malta | 3.464 | 0.002 | 0.001 | 0.005 | -0.004 | 0.001 | 0.003 | -0.004 | 0.210 | -0.389 |


| Country a | Country b | a's Fitted Growth | $\Delta \mathbf{G}$ | $\Delta \mathbf{G}^{\mathbf{m}}$ | direct | indirect | interact | $\Delta \mathbf{G}^{\mathbf{e}}$ | $\Delta \mathbf{G}^{\text {me }}$ | $\begin{gathered} \Delta \text { SSY } \\ (\%) \end{gathered}$ | $\begin{gathered} \Delta \mathbf{S S Y}^{\mathrm{m}} \\ (\%) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 |
| Italy | Switzerland | 3.464 | 0.036 | 0.008 | 0.073 | -0.059 | 0.022 | 0.040 | -0.128 | 3.187 | 10.499 |
| Jamaica | Dominican Rep. | 2.698 | -0.165 | -0.672 | 0.848 | -0.684 | -0.329 | -0.052 | 0.723 | -14.750 | -84.722 |
|  | Haiti | 2.698 | -0.165 | -1.254 | 0.847 | -0.683 | -0.329 | -0.052 | -1.268 | -14.747 | -153.034 |
| Japan | Korea | 5.361 | 0.108 | 0.119 | 0.185 | -0.149 | 0.072 | 0.139 | 0.353 | 9.645 | -4.978 |
|  | Philippines | 5.361 | 0.129 | -0.759 | 0.220 | -0.177 | 0.087 | 0.166 | -0.934 | 11.530 | -82.917 |
|  | Taiwan | 5.361 | 0.053 | -0.070 | 0.092 | -0.074 | 0.035 | 0.069 | 0.156 | 4.740 | -12.355 |
| Jordan | Israel | 1.634 | 0.016 | 0.159 | 0.636 | -0.513 | -0.107 | -0.041 | -0.356 | 1.434 | 88.681 |
|  | Syria | 1.634 | 0.062 | -0.111 | 1.010 | -0.814 | -0.133 | -0.029 | -0.294 | 5.515 | 13.099 |
| Kenya | Uganda | 1.385 | 0.184 | -0.323 | 0.426 | -0.343 | 0.102 | 0.121 | -0.492 | 16.459 | -33.066 |
| Korea | Japan | 5.788 | 0.425 | -0.308 | 0.990 | -0.798 | 0.234 | 0.459 | -0.934 | 37.952 | 75.304 |
|  | Taiwan | 5.788 | 0.096 | -0.344 | 0.265 | -0.214 | 0.044 | 0.105 | -0.068 | 8.528 | -19.555 |
| Lesotho | South Africa | 2.713 | 0.035 | -1.556 | 2.122 | -1.712 | -0.376 | -0.440 | -2.612 | 3.086 | 51.695 |
| Malawi | Zambia | 1.396 | 0.085 | -0.396 | 0.463 | -0.374 | -0.004 | 0.102 | -1.226 | 7.616 | 19.143 |
| Malaysia | Indonesia | 3.161 | 0.749 | 0.282 | 1.690 | -1.363 | 0.422 | 0.117 | -0.427 | 66.863 | -45.458 |
|  | Singapore | 3.161 | 0.037 | 0.215 | 0.126 | -0.102 | 0.013 | -0.010 | -0.226 | 3.300 | 21.953 |
|  | Thailand | 3.161 | 0.380 | -0.371 | 1.006 | -0.812 | 0.185 | 0.004 | 0.103 | 33.894 | -62.738 |
| Mali |  |  | 0.414 | 0.685 | 0.904 | -0.729 | 0.239 | 0.422 | 1.202 | 36.942 | 156.763 |
|  | Niger | 0.260 | 0.171 | 0.043 | 0.417 | -0.336 | 0.091 | 0.175 | -0.079 | 15.290 | 3.572 |
|  | Senegal | 0.260 | 0.174 | -0.192 | 0.422 | -0.341 | 0.092 | 0.177 | -0.025 | 15.501 | 19.067 |
| Malta | Algeria | 6.150 | -1.523 | -4.965 | 2.614 | -2.108 | -2.028 | -1.677 | -3.673 | -135.894 | -421.068 |
|  | Italy | 6.150 | -1.733 | -2.685 | 3.452 | -2.784 | -2.400 | -1.937 | -2.026 | -154.651 | -120.031 |
|  | Tunisia | 6.150 | -1.257 | -4.330 | 1.942 | -1.566 | -1.632 | -1.371 | -2.282 | -112.143 | -406.793 |
| Mexico | Guatemala | 1.885 | 0.057 | -0.048 | 0.084 | -0.067 | 0.041 | 0.059 | -0.038 | 5.075 | -8.159 |
|  | U.S.A | 1.885 | 0.833 | 1.038 | 1.074 | -0.866 | 0.625 | 0.865 | -0.041 | 74.304 | 204.256 |
| Netherlands |  |  | 0.051 | -0.068 | 0.373 | -0.301 | -0.021 | -0.020 | -0.002 | 4.588 | -10.388 |
|  | Germany, Fed. Rep. | 3.094 | 0.248 | -0.481 | 1.162 | -0.938 | 0.023 | 0.025 | 0.066 | 22.147 | -36.458 |
|  | United Kingdom | 3.094 | 0.234 | -1.410 | 1.119 | -0.902 | 0.018 | 0.019 | -0.191 | 20.893 | -116.219 |
| New Zealand | Australia | 2.196 | $0.204$ | $0.065$ | $1.164$ | -0.939 | $-0.021$ | $0.362$ | $0.727$ | 18.251 | 3.998 |
|  | Fiji | 2.196 | $0.009$ | -0.228 | 0.122 | -0.098 | -0.014 | 0.026 | -0.146 | 0.821 | -31.372 |
| Nicaragua | Costa Rica | 0.631 | 0.019 | 0.080 | 0.426 | -0.343 | -0.063 | 0.073 | 0.939 | 1.693 | 19.823 |
|  | Honduras | 0.631 | 0.035 | 0.211 | 0.584 | -0.471 | -0.078 | 0.109 | 0.882 | 3.108 | -2.783 |
| Niger | Algeria | 0.231 | 0.457 | 0.723 | 1.015 | -0.819 | 0.261 | 0.471 | 1.488 | 40.765 | 164.683 |


| Country a | Country b | a's Fitted Growth | $\Delta \mathbf{G}$ | $\Delta \mathbf{G}^{\mathrm{m}}$ | direct | indirect | interact | $\Delta \mathbf{G}^{\text {e }}$ | $\Delta \mathbf{G}^{\text {me }}$ | $\begin{gathered} \hline \Delta \text { SSY } \\ (\%) \\ \hline \end{gathered}$ | $\begin{gathered} \Delta \text { SSY }^{\mathbf{m}} \\ (\%) \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 |
| Niger | Benin | 0.231 | 0.136 | -0.249 | 0.351 | -0.283 | 0.068 | 0.141 | -0.273 | 12.117 | 12.430 |
|  | Mali | 0.231 | 0.228 | 0.072 | 0.562 | -0.453 | 0.120 | 0.236 | 0.234 | 20.387 | 6.741 |
| Norway | Finland | 3.044 | 0.103 | -0.357 | 0.534 | -0.430 | -0.001 | -0.018 | -0.035 | 9.152 | -35.034 |
|  | Iceland | 3.044 | 0.005 | -0.089 | 0.036 | -0.029 | -0.002 | -0.003 | -0.090 | 0.470 | -8.468 |
|  | Sweden | 3.044 | 0.162 | -0.824 | 0.757 | -0.611 | 0.015 | -0.009 | -0.783 | 14.442 | -52.154 |
| Pakistan | India | 2.728 | 1.223 | 0.853 | 1.558 | -1.257 | 0.921 | 1.213 | -0.726 | 109.106 | 92.826 |
|  | Iran | 2.728 | 0.187 | -0.494 | 0.282 | -0.228 | 0.132 | 0.185 | -0.775 | 16.674 | 30.518 |
| Panama | Colombia | 1.766 | 0.169 | 0.111 | 1.838 | -1.482 | -0.187 | 0.155 | 0.324 | 15.050 | 16.157 |
|  | Costa Rica | 1.766 | -0.018 | -0.495 | 0.535 | -0.431 | -0.121 | -0.022 | -0.323 | -1.594 | -28.217 |
| Papua New Guinea | Australia | -0.005 | 0.293 | 1.746 | 1.226 | -0.989 | 0.056 | 0.099 | 0.981 | 26.192 | 325.630 |
|  | Indonesia | -0.005 | 1.000 | 3.350 | 2.655 | -2.142 | 0.487 | 0.578 | 2.782 | 89.271 | 234.815 |
| Paraguay | Argentina | 1.362 | 0.500 | -0.363 | 1.620 | -1.307 | 0.186 | 0.809 | -1.370 | 44.581 | 94.667 |
|  | Bolivia | 1.362 | 0.159 | 0.359 | 0.718 | -0.579 | 0.020 | 0.296 | -0.348 | 14.200 | 30.436 |
|  | Brazil | 1.362 | 0.997 | 1.143 | 2.522 | -2.035 | 0.509 | 1.478 | 1.181 | 88.935 | 142.764 |
| Peru | Bolivia | 1.660 | 0.102 | 0.129 | 0.210 | -0.169 | 0.061 | 0.105 | 0.274 | 9.089 | -0.240 |
|  | Brazil | 1.660 | 0.867 | 0.898 | 1.439 | -1.160 | 0.589 | 0.886 | 2.482 | 77.403 | 69.299 |
|  | Chile | 1.660 | 0.186 | -0.188 | 0.372 | -0.300 | 0.114 | 0.191 | 0.532 | 16.617 | 0.354 |
|  | Colombia | 1.660 | 0.346 | 0.297 | 0.656 | -0.529 | 0.219 | 0.355 | 1.361 | 30.892 | 15.746 |
|  | Ecuador | 1.660 | 0.136 | 0.194 | 0.277 | -0.224 | 0.083 | 0.140 | 0.673 | 12.153 | 8.172 |
| Philippines | Indonesia | 2.301 | 0.521 | 1.021 | 0.977 | -0.788 | 0.332 | 0.550 | 1.828 | 46.470 | 50.146 |
|  | Japan | 2.301 | 0.470 | 2.301 | 0.895 | -0.722 | 0.297 | 0.497 | 2.889 | 41.957 | 286.007 |
| Portugal | Spain | 2.831 | 0.266 | 0.304 | 1.050 | -0.847 | 0.063 | 0.317 | -0.399 | 23.751 | 68.910 |
| Rwanda | Uganda | 0.890 | 0.043 | -0.395 | 0.871 | -0.703 | -0.126 | 0.461 | -0.862 | 3.815 | -27.551 |
|  | Zaire | 0.890 | 0.113 | -0.077 | 1.280 | -1.032 | -0.134 | 0.728 | -1.226 | 10.092 | -14.783 |
| Senegal | Guinea Bissau | 0.032 | 0.016 | 0.156 | 0.102 | -0.082 | -0.003 | 0.012 | 0.060 | 1.462 | 6.680 |
|  | Mali | 0.032 | 0.112 | 0.036 | 0.550 | -0.444 | 0.006 | 0.089 | -0.183 | 9.989 | -27.679 |
|  | The Gambia | 0.032 | 0.014 | 0.074 | 0.086 | -0.070 | -0.003 | 0.010 | -0.031 | 1.223 | 2.475 |
| Singapore | Hong Kong | 5.955 | -0.313 | -0.584 | 0.761 | -0.614 | -0.460 | -1.675 | -1.049 | -27.943 | -31.349 |
|  | Malaysia | 5.955 | -0.465 | -2.579 | 1.291 | -1.041 | -0.715 | -2.775 | -2.661 | -41.522 | -242.860 |
|  | Thailand | 5.955 | -0.587 | -3.269 | 2.027 | -1.635 | -0.979 | -4.214 | -2.409 | -52.401 | -343.783 |
| South Africa | Botswana | 1.138 | 0.019 | -0.038 | 0.035 | -0.028 | 0.012 | 0.012 | 0.023 | 1.666 | -5.351 |
|  | Lesotho | 1.138 | 0.025 | 0.019 | 0.047 | -0.038 | 0.016 | 0.016 | 0.037 | 2.203 | -2.360 |


| Country a | Country b | a's Fitted Growth | $\Delta \mathrm{G}$ | $\Delta \mathbf{G}^{\mathbf{m}}$ | direct | indirect | interact | $\Delta \mathbf{G}^{\text {e }}$ | $\Delta \mathrm{G}^{\mathrm{me}}$ | $\begin{gathered} \Delta \text { SSY } \\ (\%) \\ \hline \end{gathered}$ | $\begin{gathered} \Delta \mathbf{S S Y}^{\mathrm{m}} \\ (\%) \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 |
| South Africa | Swaziland | 1.138 | 0.015 | -0.008 | 0.028 | -0.023 | 0.009 | 0.009 | -0.009 | 1.331 | -1.400 |
|  | Zimbabwe | 1.138 | 0.085 | 0.064 | 0.157 | -0.127 | 0.055 | 0.054 | -0.182 | 7.585 | -3.887 |
| Spain | France | 3.156 | 0.333 | -0.312 | 0.617 | -0.498 | 0.214 | 0.393 | -0.295 | 29.710 | 13.967 |
|  | Portugal | 3.156 | 0.082 | -0.021 | 0.164 | -0.133 | 0.050 | 0.098 | 0.198 | 7.275 | -11.475 |
| Sri Lanka | Australia | 2.270 | 0.092 | -0.565 | 0.482 | -0.389 | -0.002 | 0.095 | -0.312 | 8.182 | 78.756 |
|  | India | 2.270 | 1.031 | 1.182 | 2.612 | -2.107 | 0.526 | 1.048 | 0.038 | 91.987 | 57.212 |
|  | Indonesia | 2.270 | 0.486 | 1.073 | 1.622 | -1.308 | 0.172 | 0.497 | 1.837 | 43.352 | 36.718 |
| Swaziland | South Africa | 1.088 | -0.287 | 0.042 | 2.734 | -2.205 | -0.816 | -1.132 | -1.018 | -25.599 | 59.321 |
| Sweden | Finland | 1.865 | 0.092 | 0.206 | 0.311 | -0.251 | 0.031 | 0.090 | 0.412 | 8.171 | 6.416 |
|  | Iceland | 1.865 | 0.005 | -0.023 | 0.018 | -0.015 | 0.001 | 0.005 | -0.022 | 0.429 | -2.844 |
|  | Norway | 1.865 | 0.078 | 0.354 | 0.269 | -0.217 | 0.026 | 0.076 | 0.248 | 6.975 | 22.735 |
| Switzerland | Austria | 3.486 | 0.043 | 0.238 | 0.539 | -0.435 | -0.061 | 0.092 | 0.529 | 3.856 | -8.557 |
|  | France | 3.486 | 0.264 | -0.971 | 1.518 | -1.224 | -0.029 | 0.403 | 0.968 | 23.597 | -128.333 |
|  | Germany, Fed. Rep. | 3.486 | 0.296 | -0.929 | 1.614 | -1.302 | -0.016 | 0.443 | 0.651 | 26.413 | -115.087 |
|  | Italy | 3.486 | 0.276 | -0.014 | 1.553 | -1.252 | -0.025 | 0.417 | 1.380 | 24.599 | -63.826 |
| Syria | Israel | 1.575 | 0.061 | 0.075 | 0.279 | -0.225 | 0.008 | 0.102 | -0.051 | 5.482 | 39.063 |
|  | Jordan | 1.575 | 0.043 | -0.052 | 0.202 | -0.163 | 0.004 | 0.073 | 0.004 | 3.838 | -12.032 |
|  | Turkey | 1.575 | 0.406 | 1.061 | 1.281 | -1.033 | 0.158 | 0.594 | -0.220 | 36.215 | 97.229 |
| Taiwan | Japan | 4.608 | 0.347 | 0.683 | 1.429 | -1.153 | 0.071 | 0.339 | -0.815 | 30.963 | 140.384 |
|  | Korea | 4.608 | 0.145 | 0.836 | 0.798 | -0.643 | -0.009 | 0.141 | 0.249 | 12.958 | 52.901 |
| Thailand | Malaysia | 2.573 | 0.104 | 0.217 | 0.197 | -0.159 | 0.066 | 0.091 | -0.053 | 9.317 | 30.663 |
|  | Singapore | 2.573 | 0.028 | 0.113 | 0.053 | -0.043 | 0.017 | 0.024 | -0.130 | 2.454 | 14.431 |
| The Gambia | Senegal | 2.044 | -0.398 | -1.938 | 1.576 | -1.271 | -0.703 | -0.125 | -0.755 | -35.532 | -121.751 |
| Togo | Benin | 2.597 | 0.002 | -1.794 | 0.592 | -0.477 | -0.112 | -0.070 | -1.730 | 0.205 | -83.591 |
|  | Ghana | 2.597 | 0.065 | -0.960 | 1.145 | -0.923 | -0.156 | -0.074 | -1.887 | 5.827 | -8.000 |
| Trinidad \& Tobago | Barbados | 2.034 | -0.069 | 0.027 | 0.155 | -0.125 | -0.099 | -0.030 | 0.149 | -6.135 | -10.413 |
| Tunisia | Algeria | 1.727 | 0.228 | -0.385 | 0.916 | -0.739 | 0.051 | 0.204 | -1.039 | 20.340 | -0.229 |
|  | Malta | 1.727 | 0.009 | 0.093 | 0.051 | -0.041 | -0.001 | 0.007 | -0.039 | 0.760 | 10.058 |
| Turkey | Cyprus | 2.742 | 0.014 | 0.028 | 0.025 | -0.020 | 0.009 | 0.016 | 0.036 | 1.225 | 2.993 |
|  | Greece | 2.742 | 0.089 | 0.229 | 0.159 | -0.128 | 0.058 | 0.107 | 0.408 | 7.947 | 27.006 |
|  | Iran | 2.742 | 0.250 | -0.451 | 0.426 | -0.344 | 0.168 | 0.298 | -0.759 | 22.355 | -10.525 |
|  | Syria | 2.742 | 0.073 | -0.106 | 0.130 | -0.105 | 0.048 | 0.087 | 0.182 | 6.486 | -9.886 |


| Country a | Country b | a's Fitted Growth | $\Delta \mathbf{G}$ | $\Delta \mathbf{G}^{\mathbf{m}}$ | direct | indirect | interact | $\Delta \mathrm{G}^{\text {e }}$ | $\Delta \mathbf{G}^{\text {me }}$ | $\begin{gathered} \Delta \text { SSY } \\ (\%) \\ \hline \end{gathered}$ | $\begin{gathered} \hline \Delta \text { SSY }^{\mathbf{m}} \\ (\%) \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 |
| U.S.A | Canada | 2.995 | 0.069 | 0.069 | 0.070 | -0.056 | 0.055 | 0.060 | 0.221 | 6.137 | 3.764 |
|  | Mexico | 2.995 | 0.162 | -0.072 | 0.163 | -0.132 | 0.131 | 0.141 | 0.203 | 14.469 | -19.768 |
| Uganda | Kenya | 0.426 | 0.203 | 0.636 | 0.580 | -0.468 | 0.091 | 0.370 | 0.938 | 18.092 | 62.210 |
|  | Rwanda | 0.426 | 0.080 | 0.069 | 0.252 | -0.203 | 0.031 | 0.153 | 0.498 | 7.145 | 3.105 |
|  | Zaire | 0.426 | 0.298 | 0.379 | 0.803 | -0.647 | 0.143 | 0.530 | 0.051 | 26.588 | 20.015 |
| United Kingdom | France | 1.308 | 0.209 | 0.702 | 0.444 | -0.358 | 0.123 | 0.211 | 0.509 | 18.670 | 55.590 |
|  | Ireland | 1.308 | 0.016 | 0.042 | 0.038 | -0.030 | 0.009 | 0.016 | 0.025 | 1.447 | 1.118 |
|  | Netherlands | 1.308 | 0.064 | 0.376 | 0.145 | -0.117 | 0.036 | 0.065 | 0.035 | 5.731 | 31.593 |
| Uruguay | Argentina | 0.639 | 0.311 | 0.341 | 1.563 | -1.261 | 0.008 | 0.824 | -0.129 | 27.727 | 40.951 |
|  | Brazil | 0.639 | 0.702 | 1.820 | 2.462 | -1.986 | 0.226 | 1.511 | 2.417 | 62.674 | 86.575 |
| Venezuela | Barbados | 0.393 | 0.013 | 0.036 | 0.042 | -0.034 | 0.005 | 0.015 | 0.049 | 1.126 | 1.382 |
|  | Brazil | 0.393 | 0.700 | 1.980 | 1.565 | -1.262 | 0.397 | 0.806 | 3.021 | 62.479 | 70.930 |
|  | Colombia | 0.393 | 0.276 | 1.021 | 0.749 | -0.604 | 0.132 | 0.327 | 1.745 | 24.669 | 21.418 |
|  | Guyana | 0.393 | 0.020 | 0.011 | 0.066 | -0.053 | 0.007 | 0.025 | -0.137 | 1.789 | -4.311 |
| Yugoslavia | Greece | 4.050 | 0.073 | -0.124 | 0.247 | -0.199 | 0.025 | 0.109 | 0.311 | 6.508 | -8.321 |
| Zaire | Central Afr. Rep. | 0.810 | 0.047 | -0.023 | 0.078 | -0.063 | 0.032 | 0.043 | 0.004 | 4.162 | 1.888 |
|  | Congo | 0.810 | 0.035 | -0.025 | 0.059 | -0.048 | 0.024 | 0.032 | 0.127 | 3.150 | 4.830 |
|  | Rwanda | 0.810 | 0.081 | 0.003 | 0.135 | -0.109 | 0.055 | 0.075 | 0.325 | 7.269 | 1.737 |
|  | Uganda | 0.810 | 0.171 | -0.005 | 0.277 | -0.224 | 0.117 | 0.157 | 0.241 | 15.267 | 5.880 |
|  | Zambia | 0.810 | 0.089 | 0.097 | 0.148 | -0.119 | 0.061 | 0.082 | -0.076 | 7.963 | 23.515 |
| Zambia | Malawi | 1.000 | 0.178 | -0.001 | 0.521 | -0.420 | 0.077 | -0.008 | 0.731 | 15.904 | -38.779 |
|  | Zaire | 1.000 | 0.496 | -0.093 | 1.215 | -0.980 | 0.261 | 0.063 | 0.273 | 44.304 | -61.472 |
|  | Zimbabwe | 1.000 | 0.197 | 0.175 | 0.567 | -0.458 | 0.087 | -0.006 | 0.672 | 17.541 | 16.938 |
| Zimbabwe | Botswana | 1.316 | 0.028 | -0.157 | 0.107 | -0.087 | 0.008 | 0.017 | 0.251 | 2.543 | -19.579 |
|  | South Africa | 1.316 | 0.424 | -0.114 | 1.158 | -0.934 | 0.200 | 0.298 | 0.687 | 37.870 | 59.791 |
|  | Zambia | 1.316 | 0.127 | -0.140 | 0.429 | -0.346 | 0.044 | 0.080 | -0.755 | 11.351 | -13.648 |

Table 9 - Pairs of countries that would both have gained from full political integration ( $\Delta \mathbf{G}^{\mathrm{m}}>0$ for both countries)

| Argentina | Chile |
| :--- | :--- |
| Bolivia | Brazil |
| Brazil | Colombia |
| Brazil | Guyana |
| Brazil | Paraguay |
| Brazil | Peru |
| Canada | U.S.A |
| Colombia | Peru |
| Denmark | Federal Republic of Germany |
| France | Federal Republic of Germany |
| India | Pakistan |
| India | Sri Lanka |
| Malaysia |  |
| Indonesia | Niger |
| Mali |  |





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[^1]:    ${ }^{1}$ The methodology can be easily extended to a case in which more than two countries are considering integrating politically. We do not pursue such an extension in this paper.
    ${ }^{2}$ This literature is not directly concerned with the effects of national borders on economic growth. In his important book on border effects in international trade, Helliwell [1998, chapter 6, p. 112] states that "assessing the possible growth implications of home preferences is not a job for a one-handed economist, nor for the faint of heart". We are not aware of research that tries to accomplish such measurement.

[^2]:    ${ }^{3}$ The classical reference is Viner (1950). For a textbook exposition see, for example, Vousden [1990], ch. 10.
    ${ }^{4}$ For a survey of the regional-integration literature, see Baldwin and Venables [1995].
    ${ }^{5}$ The political economy of trade barriers and protectionism is surveyed, for instance, by Rodrik (1995).
    ${ }^{6}$ The effect of the extent of the domestic market on growth has also been investigated by Ades and Glaeser [1999]. A recent confirmation of the Alesina-Spolaore-Wacziarg hypothesis on the relationship between size, openness and growth has been provided by Alcalá and Ciccone (2003).

[^3]:    ${ }^{7}$ It is worth noting that in our model the effect of market size on productivity will not be due to a technology with increasing returns. In fact we will use a production function with constant returns to scale.
    ${ }^{8}$ As usual, the results generalize to any standard CRRA utility function $\left(C_{i t}^{1-\sigma}-1\right) /(1-\sigma)$ with $\sigma>0$.

[^4]:    ${ }^{9}$ Equation (8) is obtained from equation (3) by susbtituting domestically-produced and imported intermediate inputs with their equilibrium values, as specified in equations (51) and (52), and $K_{i}$ with its stead-state value in (7).
    ${ }^{10}$ For a derivation of these standard results see, for example, Barro and Sala-i-Martin (1995).

[^5]:    ${ }^{11}$ In this model we abstract from international borrowing and lending - hence exports are always equal to imports in equilibrium. Therefore, measuring openness as exports/output is identical, up to a scalar multiplication, to measuring openness as (exports + imports)/output.

[^6]:    ${ }^{12}$ For simplicity we assume that policy-makers have measure zero in the economy, and therefore their rents and costs do not affect per capita consumption, capital accumulation and production directly, but only through policy decisions.
    ${ }^{13}$ A specification in which policy-makers attach weight both to their rents (contributions from lobbies) and to their citizens' welfare is provided, for instance, in Grossman and Helpman's [1994] classic analysis of protectionism. For a survey of this extensive literature see, for example, Rodrik [1995]. In our model we do not model a private demand for protection explictly, but just assume that the government's rents are a function of barriers. Our main results would go through even if $\psi_{n}=0$. More generally, we will assume $0 \leq \psi_{n}<1$.

[^7]:    ${ }^{14}$ We assume that each country will reduce barriers taking other countries' barriers as given (Nash equilibrium). That is, in each first-order condition, other countries' barriers will be taken as given at their equilibrium level (i.e., $\xi_{n}^{s s}=\frac{\xi}{2}-\lambda_{n}^{s s}$ for all $n \neq a$ ). By contrast, joint maximization of world welfare would imply lower barriers.
    ${ }^{15}$ The result can be generalized to the case of three or more countries.
    ${ }^{16}$ It is immediate to check that (23) reduces to (25) for $\alpha=1 / 2$.

[^8]:    ${ }^{17}$ On the costs of larger, more heterogeneous countries see Alesina and Spolaore [1997, 2003].
    ${ }^{18}$ For instance, a previously landlocked country can gain easier access to the sea as a result of a political

[^9]:    ${ }^{19}$ Below we will examine how to account for changes in the $Z$ and $W$ variables under a merger.
    ${ }^{20}$ In section 3.3 below, we discuss an alternative method that allows us to include the error term component of the growth effect of mergers, using the estimated values of error term in the original countries.

[^10]:    ${ }^{21}$ For example, we could assume that the merged country is assigned the best - or worst - values of the $Z$ and $W$ variables from each of country $a$ and $b$. We choose an intermediate - and more reasonable - assumption by assigning to the merged country the population-weighted average of these variables from countries $a$ and $b$.
    ${ }^{22}$ Of course, in the case of the land area, the merged variable is the sum of the corresponding areas of countries $a$ and $b$. For the dummy variables in our specification, the definitions of the merged variables are as follows: the merged country is landlocked if both $a$ and $b$ are landlocked; the merged country is an island if $b o t h ~ a$ and $b$ are islands; the merged country is an oil exporter if either $a$ or $b$ is an oil exporter.

[^11]:    ${ }^{23}$ We can further decompose the steady-state determination effect into the term $\gamma_{1} \log \frac{y_{m t-\tau}}{y_{a t-\tau}}$ which reflects differences in initial income and the terms that are functions of $\left(W_{m t}-W_{a t}\right)$ and $\left(Z_{m t}-Z_{a t}\right)$, which reflect differences in steady-state determinants proper. For identical values of the $Z$ and $W$ variables, if country $a$ starts out with an initial income that is lower than country $b$ 's, full integration will slow $a$ 's growth simply because it will raise its initial income - the force of convergence implies that countries grow slower, the closer they are to their steady-states.
    ${ }^{24}$ This is not the case when we do not take into account the steady-state determination effect (section 3.1), because post-merger $Z$ and $W$ variables still differ across $a$ and $b$.

[^12]:    ${ }^{25}$ The only reason $\nu_{a t}$ remains in this equation is the nonlinearity of the effect of country size on growth brought forth by the interaction term between openness and size in the growth equation.

[^13]:    ${ }^{26}$ See Barro and Sala-i-Martin (1995), p. 37 and p. 82 for a derivation of this standard specification in the context of the neoclassical growth model.

[^14]:    ${ }^{27}$ Note that $\beta_{1}$, the conditional convergence coefficient, is negative.

[^15]:    ${ }^{28}$ See http://webhost.bridgew.edu/baten/.
    ${ }^{29}$ In the case of Germany, this is due to reunification in 1989. The new version of the PWT only features data for reunified Germany since 1990. Our estimates of merger effects refer to West Germany prior to 1990.

[^16]:    ${ }^{30}$ Following Alesina, Spolaore and Wacziarg [2000], these are dummy variables for small countries, small islands, and the interaction terms between population and the each of dummy variables for small countries, small islands, islands, and landlocked countries. As long as they are jointly excludable from the growth regression, geographic variables such as these are likely to be plausibly exogenous with respect to growth, yet affect the level of openness. See Frankel and Romer [1999] for a further details on employing geographic variables to instrument for openness in growth regressions, and for arguments that these variables are indeed excludable from a growth specification when other determinants of growth are controlled for.
    ${ }^{31}$ See Wacziarg [2001] for further technical details on the use of 3SLS to estimate systems of equations in a cross-country growth context.

[^17]:    ${ }^{32}$ These test statistics are available upon request. Similar tests appeared in Alesina, Spolaore and Wacziarg [2000], who also concluded that the geographic determinants of openness were excludable from the growth

[^18]:    ${ }^{34}$ For further details on panel-3SLS estimators, see Wacziarg [2001].
    ${ }^{35}$ We also considered mergers between proximate islands and up to five neighboring countries, such as the United Kingdom and Ireland, or the United Kingdom and France. Our results pertain to a total of 123 hypothetical mergers of country pairs (i.e. 246 merger experiments).

[^19]:    ${ }^{36}$ This is another way of saying that Italy was a faster growing country than France over the time period covered in the sample. In fact, the average observed annual growth rate of per capita income in Italy over the 1960-1989 period was $3.40 \%$, while for France it was $2.94 \%$. Our model predicts that, if France and Italy had merged, their unified growth rate over this period would have been $3.15 \%$ per year (under "full integration").

[^20]:    ${ }^{37}$ In general, accounting for the residual effect has a much smaller effect on estimates of pure size mergers than it does on estimates of full integration, because the former only involves the residual from the openness regression (multiplied by the coefficient on openness in the growth regression), while the latter involves the population weighted average of the residual from the growth regression. See equations (40) and (41).

[^21]:    ${ }^{38}$ All of these pairs are also composed of countries that would both have benefited from size mergers with each other.

